Elevated moral condemnation of third-party violations in multiple sclerosis patients

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**Elevated moral condemnation of third-party violations
in multiple sclerosis patients**

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Keywords: morality, punishment, empathy, alexithymia, multiple sclerosis, moral luck

Abbreviations: BF₁₀ : Bayes Factor in favour of the alternative hypothesis; CI : confidence interval; DDF : difficulty describing feelings; DIF : difficulty identifying feelings; EC : empathic concern; EDSS : Expanded Disability Status Scale; EOT : externally-oriented thinking; F : fantasy; HC : healthy controls; MS : multiple sclerosis; MSSS : Multiple Sclerosis Severity Score; *n* : sample size, PT : perspective taking; PD : personal distress; TAS-20 : Toronto Alexithymia Scale total score.

Abstract

Recent research has demonstrated impairments in social cognition associated with multiple sclerosis (MS). The present work asks whether these impairments are associated with atypical moral judgment. Specifically, we assessed whether MS patients are able to integrate information about intentions and outcomes for moral judgment (i.e., appropriateness and punishment judgments) in the case of third-party acts. We found a complex pattern of moral judgments in MS patients: although their moral judgments were comparable to controls' for specific types of acts (e.g., accidental or intentional harms), they nevertheless judged behaviors to be less appropriate and endorsed more severe punishment across the board, and they were also more likely to report that others' responses would be congruent with theirs. Further analyses suggested that elevated levels of externally oriented cognition in MS (due to co-occurring alexithymia) explain these effects. Additionally, we found that the distinction between appropriateness and punishment judgments, whereby harmful outcomes influence punishment judgments to a greater extent than appropriateness judgments, was preserved in MS despite the observed disruptions in affective and motivational components of empathy. The current results inform the two-process model for intent-based moral judgments as well as possible strategies for improving the quality of life in MS patients.

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3 **1. Introduction**
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5 MS is an autoimmune disease involving demyelination of nerve cells as well as neural
6 inflammation and degeneration, which leads to a wide range of physical, cognitive, and
7 psychiatric signs and symptoms. In addition to the study of physical impairments and
8 neuropsychological dysfunction (Feinstein et al., 2014; Pepping et al., 2013; Rocca et al.,
9 2015), recent research has also begun to unravel problems associated with social cognition in
10 MS.
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19 **1.1 MS and deficits in social cognition**
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22 Social cognition is an umbrella term for the set of processes necessary for effectively dealing
23 with social situations, including perception and evaluation of socio-emotional stimuli as well
24 as regulation and modulation of one’s social behavior according to such evaluations
25 (Adolphs, 2009).
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32 Studies in MS have shown impairments in theory of mind (ToM), the capacity to represent
33 others’ epistemic states (i.e., thoughts, beliefs, desires, etc.; Koster-Hale & Saxe, 2013). In
34 particular, MS patients perform poorly in extracting information about intentionality,
35 representing first-and second-order false beliefs, and detecting social faux pas in verbal/non-
36 verbal ToM assessment tasks (Banati et al., 2010; Charvet et al., 2014; Henry et al., 2011;
37 Pöttgen et al., 2013; Roca et al., 2014). These deficits remain even after controlling for the
38 co-occurring neuropsychological impairments, such as executive function deficits (Charvet et
39 al., 2014; Pöttgen et al., 2013). Importantly, ToM deficits appear from the very early stage of
40 the disease in young adults (Kraemer et al., 2013) and are found even in pediatric-onset MS
41 patients (Charvet et al., 2014). These ToM deficits are also known to get worse with the
42 disease duration/physical disability (Banati et al., 2010) and cognitive impairment (Ouellet et
43 al., 2010) in MS.
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MS patients have also been shown to perform atypically on empathy tasks that index the cognitive ability to form representations of feeling states (pain, emotions, etc.) in others and affective sharing of these states (de Vignemont & Singer, 2006). Self-report questionnaires reveal disruptions in general empathizing skills in MS (Kraemer et al., 2013), with a more specific reduction in empathic concern alongside elevated levels of personal distress (Gleichgerricht et al., 2015). MS patients also show impaired performance on tasks requiring them to infer emotional states from facial (Henry et al., 2009), bodily (Cecchetto et al., 2014), or prosodic (Beatty et al., 2003) affective cues, as well as an abnormal pattern of neural processing of emotional faces (Jehna et al., 2011; Krause et al., 2009).

Additionally, MS patients have been shown to exhibit elevated levels of alexithymia, which is a dimensional personality construct defined by difficulties in analyzing, identifying, and communicating emotional states and externally-focused cognitive style (Nemiah et al., 1976). Whereas the preponderance rate of clinical alexithymia is about 10% in the healthy population (Bird & Cook, 2013), it is found to vary from 13.8 to as high as 50% in the MS population. Alexithymia also reliably predicts elevated levels of fatigue, depression, and anxiety in MS (Bodini et al., 2008; Chahraoui et al., 2014; Gay et al., 2010; Gleichgerricht et al., 2015).

1.2 *Intent-based moral judgments and its neural basis*

The two-process model of intent-based moral judgments (Cushman, Sheketoff, Wharton, & Carey, 2013; Cushman, 2008, 2015) posits two distinct processes that guide moral judgment: (i) a causal reasoning process, active in the presence of a harmful outcome (e.g., victim suffering), that involves an analysis of the agent's causal role in producing the outcome ("causal responsibility = bad"); (ii) an intent-based reasoning process that involves an analysis of the agent's intent to bring about the outcome, leading to condemnation in the

presence of a culpable mental state (“malicious belief/desire/intent = bad”). Although the outputs of these processes need not conflict (i.e., in the case of neutral acts or intentional harms), these two processes may yield different judgments (i.e., in case of accidental or attempted harms). The ultimate judgment in the case of conflicting outputs is the result of a competitive interaction between antecedent evaluations, and informed by the relative weight (which itself is determined, in part, by underlying personality traits; Prehn et al., 2008) assigned to the output of each process (Buckholtz et al., 2015; Young, Cushman, Hauser, & Saxe, 2007).

Mental-state reasoning process: Abundant evidence shows that both older children and healthy adults rely primarily on mental state information when evaluating third-party harms (Alter, Kernochan, & Darley, 2007; Baird & Astington, 2004; Barrett et al., 2016; Cushman, 2008; Gummerum & Chu, 2014). Behaviorally, individuals tend to forgive accidental harms based on innocent intentions, while they condemn attempted harms based on malicious intentions despite the absence of harmful outcomes (Cushman, 2008). At the neural level, the rTPJ has been shown to be a key region within the ToM network for mental state attribution during moral judgment (for a review, see Young & Tsoi, 2013). The rTPJ is recruited most robustly in response to attempted harms, when the perpetrator intends but fails to harm someone, and thus the condemnation relies heavily on intent information (Gan et al., 2015; Young et al., 2007; Young & Saxe, 2008); in fact, disrupting activity in the rTPJ leads to reduced condemnation of attempted harms (Young, Camprodon, Hauser, Pascual-Leone, & Saxe, 2010). Additionally, neurological patients (Baez et al., 2014a, 2014b, 2016a, 2016b; Ciaramelli, Braghittoni, & di Pellegrino, 2012; Young, Bechara, et al., 2010) and sadistic individuals (Trémolière & Djeriouat, 2016) deliver more favorable assessments of attempted harms due to a reduced emotional response to harmful intent. Note that in case of condemning attempted harms the causal reasoning process remains silent in the absence of a

harmful outcome, and the intent-based process dominates moral judgment (Cushman, 2008; Young et al., 2007).

Compared to attempted harms, accidental harms pose a more difficult case. Forgiving accidental harms requires a robust representation of innocent intent that can counteract a prepotent tendency to condemn the actor based on the harmful outcome (cf. Moran et al., 2011). Accordingly, individuals with a higher overall response in the rTPJ (Young & Saxe, 2009) and greater differentiation of intentional versus accidental harms in the spatial pattern of activity in the rTPJ (Chakroff et al., 2016; Koster-Hale, Saxe, Dungan, & Young, 2013) tend to forgive accidental harms more. This can happen as early as 62 ms post-stimulus (Decety & Cacioppo, 2012) and involves down-regulation of emotional arousal encoded in the amygdala in response to harm (Hesse et al., 2016; Ngo et al., 2014; Treadway et al., 2014; Yu, Li, & Zhou, 2015). What is more, stimulating this patch of cortex increases the role that belief information plays in moral judgments (Ye et al., 2015), as evidenced by more lenient judgments of accidental harms (Sellaro et al., 2015). Meanwhile, imposing a cognitive load on participant and thereby taxing integration of mental state information, leads to the opposite pattern (Buon, Jacob, Loissel, & Dupoux, 2013a). Finally, populations with impaired mental state inference, like autism spectrum disorder, show abnormal patterns in the rTPJ (i.e., lack of differentiation between intentional and accidental harms in the spatial patterns of activity in the rTPJ; Chakroff et al., 2016; Koster-Hale et al., 2013) and deliver harsher moral judgments of accidental harms (Buon, Dupoux, et al., 2013b; Moran et al., 2011; Salvano-Pardieu et al., 2015; but see Baez et al., 2012).

Causal reasoning process: Causal analyses of harmful events begin with the detection of a harmful outcome and subsequent search for a causally responsible agent (Sloman, Fernbach, & Ewing, 2009). The degree to which individuals pay attention to the causal role of actors who accidentally bring about negative outcomes is in turn determined by the extent to which

they empathize with the victim. In other words, understanding and feeling victim distress can motivate individuals to condemn accidental harm-doers more based on causal involvement. Evidence on individual differences in moral judgment is broadly consistent with this idea. For example, individuals who score high on self-report measures of dispositional empathy are more inclined to condemn accidental harms (Trémolière & Djeriouat, 2016, Study 1). In addition, individuals with a certain genetic variation of the oxytocin receptor gene that predisposes them to being more empathic are more reluctant to exculpate accidental harm-doers (Walter et al., 2012). Subclinical (e.g., alexithymia; Patil & Silani, 2014) and clinical (e.g., psychopathy; Young, Koenigs, Kruepke, & Newman, 2012) personalities characterized by reduced empathic concern for others also exhibit an increased tendency to forgive accidents, arguably because they are less motivated to hold the agent causally responsible in the absence of strong empathic aversion. Finally, while evaluating harmful situations, participants in general spend more time looking at the victim than the perpetrator and exhibit increased activity in the empathy network (Decety, Michalska, & Kinzler, 2012).

Differential effect of moral luck on punishment and wrongness: Importantly, according to the two-process model, information about intent and cause are recruited to different degrees for different types of moral judgments. In particular, blame/punishment judgments rely to a greater degree on outcome information as compared to wrongness/appropriateness judgments, which rely more on intent information (Cushman, 2008, 2015; Cushman et al., 2009; Cushman et al., 2013). For instance, although the behaviors of two drivers who fall asleep at the wheel while driving under the influence are judged to be equally *wrong* or *inappropriate*, the driver who runs over and kills someone is *punished* more severely than the driver who runs into a tree. This asymmetric reliance on outcomes for punishment judgments (vis-à-vis appropriateness) has convincingly been argued to be an upshot of the ultimate evolutionary function of punishment (Martin & Cushman, 2016), i.e., to utilize the learning

capacity of social partners to modify harmful behavior, including unintentionally harmful behavior. At the mechanistic level, however, this approach is implemented via inflexible moral outrage towards the harm-doer (Martin & Cushman, 2016), possibly stemming from empathy with the victim.

To summarize, judgments of third-party moral acts rely on ToM skills for generating robust representations of the agent's mental states, irrespective of whether a harmful outcome is present. Empathic evaluation of victim suffering (in the presence of a harmful outcome) additionally contributes to moral evaluation. Additionally, this empathic aversion exerts greater influence on punishment judgments versus wrongness judgments. Finally, these underlying psychological mechanisms are neurally implemented in interactions between limbic and frontotemporal structures.

1.3 Past work and current study

Given the aforementioned sociocognitive and socioaffective deficits in the MS population, particularly in ToM and empathy skills, we expect MS patients to exhibit irregular patterns of moral evaluations that rely on these very processes, as compared to the neurotypical population. In addition to these deficits, prior work on neuroanatomical and neurofunctional correlates of MS reveals that the disease presents with: (a) frontotemporal changes such as atrophy (Calabrese et al., 2010, Bonavita et al., 2013); (b) affected structural (Fox, 2008) and (c) functional (Sacco et al., 2013) frontotemporal connections; and (d) aberrant amygdala-frontotemporal activation at the resting state (Nigro et al., 2015), and also during the processing of salient affective stimuli (Passamonti et al., 2009, Sacco et al., 2013). Given that research with neurological populations with involvement of frontotemporal regions have revealed deficits in intent-based moral judgments (Baez et al., 2014a, 2014b, 2016a, 2016b; Ciaramelli et al., 2012; Young et al., 2010), we expected a similar disruption in MS. Thus far,

only one study has investigated moral cognition in the MS population (Gleichgerricht et al., 2015). MS patients and healthy controls were presented with moral dilemmas asking about the permissibility of sacrificing the welfare of the few in favour of the aggregate welfare; MS patients judged the act of sacrificing the few for the greater good (utilitarian response) as less morally permissible due to increased emotional reactivity. This is a surprising result given that MS patients showed elevated levels of alexithymia and reduced empathic concern, both of which are associated with increased utilitarian inclinations (Gleichgerricht & Young, 2013; Koven, 2011; Patil & Silani, 2014a; Patil et al., 2016). Additionally, MS patients also showed elevated egocentric moral attitudes (compared to healthy controls) in that they reported that others would respond the same way as they did.

In the current exploratory study, we extended this work with a task that featured relatively ordinary and familiar scenario settings (as compared to more contrived moral dilemma contexts). Furthermore, the task featured not only intentional harms, but also unintentional or accidental harms, and behaviors that were performed without the goal of maximizing aggregate welfare. This task may therefore probe the subtle effects of social cognition deficits on moral cognition in MS patients. In the light of the prevalent ToM deficits, on the one hand, we expected MS patients to provide lenient judgments of attempted harms or harsh judgments of accidental harms. On the other hand, empathic deficits were expected to lead to lenient judgments of accidental harms. In light of the complicated pattern of social cognitive deficits in MS with divergent associated predictions, we did not have more specific predictions and relied on exploratory data analysis

2. Materials and Methods

2.1 Participants

Thirty-eight consecutive patients (86.8% female) who fulfilled McDonald criteria (Polman et al., 2011) for Relapsing-Remitting Multiple Sclerosis (RRMS) were recruited for the present study. A large portion of these patients were previously recruited for a study on moral judgment (Gleichgerricht et al., 2015). Patients reported no history of alcohol/drug abuse, major psychiatric disorder, or traumatic brain injury. They were assessed at least 90 days after the most recent relapse episode and had been off steroid treatment for at least three months. Patients who were having a relapse in their MS were not included in the study. All the patients were receiving disease-modifying therapies at the time of assessment and scored above the proposed cut-off score for the Mini-Mental State Examination (MMSE, Folstein et al., 1975; score > 24). Their mean Expanded Disability Status Scale (EDSS) score (Kurtzke, 1983) was 1.66 (SD = 1.6, range = 0–6, median = 1.25), indicating mild MS. Mean disease duration (in number of years) was 10.60 (SD = 8.7, range = 1.38 – 39.3, median = 9.01). Mean number of relapses was 3.4 (SD = 1.92, range = 2–12, median = 3), and mean Multiple Sclerosis Severity Score (MSSS) (Roxburgh et al., 2005) was 2.35 (SD = 2.4, range = 0.01–8.64, median = 1.29) points.

Thirty-eight age-, gender- and level of education-matched volunteers were also included in the healthy controls (HC) group after having undergone screening to ensure absence of history of drug abuse, neurological or neuropsychiatric disorders. Neither MS patients nor healthy controls were financially compensated and voluntarily participated in the study. All participants signed an informed consent form before participating in this study, which was approved by the Ethics Committee at the Institute of Cognitive Neurology (INECO, Buenos Aires, Argentina).

2.2 Empathy and alexithymia measurement

The Interpersonal Reactivity Inventory (IRI) (Davis, 1983; Spanish-validated version: Pérez-Albéniz et al., 2003) was used to assess specific aspects [fantasizing, empathic concern (EC), perspective-taking (PT), and personal distress (PD)] of dispositional empathy. Inspired by recent theoretical and empirical work involving empathy and moral cognition (Decety & Cowell, 2014; Decety & Yoder, 2015), we used EC, PT, and PD subscales to denote motivational, cognitive, and affective components of empathy, respectively. Moreover, based on recent psychometric assessments of the IRI questionnaire (Baldner & McGinley, 2014), we decided *a priori* not to explore the fantasy subscale beyond descriptive statistics, as it does not map well onto the current theorizations of empathy.

To assess levels of trait alexithymia, we used the validated Spanish version of the Toronto Alexithymia Scale-20 (TAS-20) questionnaire (Bagby, Taylor, & Parker, 1994; Spanish version: Martínez Sánchez, 1996) consisting of three subscales: Difficulty Describing Feelings (DDF, 5 items, e.g., “I am often puzzled by sensations in my body.”), Difficulty Identifying Feelings (DIF, 7 items, e.g., “I often don’t know why I am angry.”), and Externally-Oriented Thinking (EOT, 8 items, e.g., “I prefer to analyze problems rather than just describe them”). Out of these three constructs, DDF and DIF together represent affective aspects of alexithymia (Zackheim, 2007), while EOT is construed as a cognitive/attentional component of alexithymia as it is less affective in its scope (Moriguchi & Komaki, 2013). Following the criteria proposed by the original authors (Bagby et al., 1994), individual scores were used to classify participants as either non-alexithymic (score ≤ 51), borderline alexithymic (scores of 52 to 60), or as alexithymic (scores ≥ 61).

For more details about these questionnaires and their internal reliability analysis, see Supplementary Text S1.

2.4 Moral judgment task

Experimental stimuli consisted of 24 unique stories, equally divided among four conditions resulting from a 2 (belief: neutral, negative) $\times$ 2 (outcome: neutral, negative) within-subject design such that agents in the scenario produced either a neutral outcome or a harmful outcome while acting with the belief that they were causing either a neutral outcome or a harmful outcome (see Figure 1). The magnitude of harm severity varied freely across scenarios from mild, to severe, to fatal injuries. We note that the stories were not balanced across conditions, i.e. a given story appeared only in one condition (e.g. accidental), but never in any other condition. Stories were matched for word length across conditions. All scenarios were adapted in Spanish from Young et al. (2010) (see Supplementary Text S2 to see the scenario-by-condition breakdown and labels for scenarios from original material).

Each scenario was presented on screen for as long as the patient needed in order to make sure that the subjective feeling of time pressure did not affect responses and also to alleviate working memory demands for patients. Each story consisted of four cumulative segments: (i) *background*: this segment provided the setting in which the story took place; (ii) *foreshadow*: this segment foreshadowed whether the outcome would be neutral or harmful; (iii) *belief*: this segment provided information about whether the agent was acting with a neutral or harmful belief; (iv) *consequence*: this final segment revealed the outcome of the agent's action.

We provide below an example of one scenario ("Vitamin") that appeared in the accidental harm condition:

Juan is instructed by a doctor to give his senile wife pills for her heart disease. The doctor says that she must not intake vitamin K within an hour to take the pills safely. One day, his wife tries a new kind of fruit. The new kind of fruit is high in vitamin K, so it is deadly for Juan's wife to take the pills right away. Juan does his research and believes that the new kind of fruit does not have vitamin K and that it is safe to give her the pills. Juan gives his wife the pills right away. His wife dies of heart failure.

The order of presentation of scenarios was randomized across subjects. After reading each scenario, participants provided three types of moral judgments which always appeared in the same order:

(i) *appropriateness*: "How appropriate was it for [the agent] to do [nature of the action]?" (7-point Likert scale: 1 = completely inappropriate; 7 = completely appropriate);

(ii) *punishment*: "How severely should [the agent] be punished for [nature of the action]?" (7-point Likert scale: 1 = no punishment, 7 = severe punishment);

(iii) *egocentrism*: "Out of 100 people answering to this scenario, how many do you think would answer like you?" (continuous scale: from 0 to 100).

Appropriateness ratings were later reverse-scored to be positively associated with the punishment judgments, and higher scores thus reflect perceived *inappropriateness* of the agent's behavior. Participants had up to 30 s to respond to each question.

2.5 Statistical analysis

Data was analyzed using SPSS 22 (for classical inference) and JASP 0.7.4 (for Bayesian inference). Effect size measures are reported (Lakens, 2013, 2015a) along with Bayes factors (cf. recommendations by Lakens, 2015b), which are provided in supplementary materials. We follow guidelines provided in Nimon (2012) to ensure that our data met the statistical assumptions associated with the general linear model-based statistical tests that we employed. Correlation analysis was carried out using Spearman's *rho* since it is more robust to univariate outliers than Pearson's *r* (Pernet et al., 2012). Welch's *t*-test was used as a default for between-group comparisons instead of Student's *t*-test because it accounts for unequal variances between groups (cf. Lakens, 2015c). As recommended (Weissgerber et al., 2015),

in addition to bar graphs in the main text, we have also provided univariate scatter-plots for main dependent variables in the supplementary material.

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3. Results

3.1 Demographic information

The two groups were well-controlled for demographic variables as they did not differ in terms of their age (HC = 39.3 (8.1), MS = 42.3 (11.3); $t(66.69) = 1.309$, $p = 0.195$, $BF_{10} = 0.497$), number of years of formal education (HC = 15.7 (1.8), MS = 15.4 (2.8); $t(64.16) = -0.680$, $p = 0.499$, $BF_{10} = 0.290$), or gender composition (proportion of females; HC = 81.60%, MS = 86.80%; $\chi^2(1) = 0.396$, $p = 0.529$, $BF_{10} = 0.367$).

3.2 Group differences in Alexithymia and Empathy

As expected from prior work (Gleichgerricht et al., 2015), we found that patients with MS showed elevated levels of all aspects of alexithymia (Table 1). There was a significant difference between individuals in HC and MS groups ($\chi^2(2) = 24.346$, $p < 0.001$, $BF_{10} = 41975$) in the composition of alexithymic (HC = 1/38 vs. MS = 12/38), borderline alexithymic (HC = 1/38 vs. MS = 10/38), and non-clinically alexithymic (HC = 36/38 vs. MS = 16/38). Moreover, the MS population showed reduced scores on the motivational component of empathy (EC), increased levels of affective sharing (PD), and did not differ on the cognitive component of empathy (PT) relative to the HC group (see Table 1).

[Table 1 about here]

We ensured that these differences were not due to poor internal reliability of measures in MS population by checking Cronbach’s alpha values for each subscale and for each group (see Supplementary Text S1).

3.3 Descriptive statistics for moral judgments

Before averaging judgments across scenarios in each condition at the subject-level, we ascertained that the items showed good internal reliability (see Supplementary Text S3).

Average ratings for all 24 scenarios for each group are provided in Supplementary Text S4.

Descriptive statistics for appropriateness, punishment, and relative moral judgments averaged at the group-level are reported in Table 2.

[Table 2 about here]

3.4 Group differences in moral judgments about appropriateness

To assess whether the two groups relied to a different degree on belief and outcome information while evaluating the moral appropriateness of the actions, we conducted a $2(\text{belief}) \times 2(\text{outcome}) \times 2(\text{group})$ mixed ANOVA. This analysis revealed a main effect of belief ($F(1,74) = 1120.59, p < 0.001, p\eta^2 = 0.938, \omega^2 = 0.935$) and outcome ($F(1,74) = 4.838, p = 0.031, p\eta^2 = 0.061, \omega^2 = 0.048$) but no group-by-belief ($F(1,74) = 2.029, p = 0.173, p\eta^2 = 0.025, \omega^2 = 0.001$) or group-by-outcome ($F(1,74) = 0.481, p = 0.490, p\eta^2 = 0.006, \omega^2 \sim 0$) interactions. In other words, the behavior of agents who were acting with a harmful intent was judged to be more inappropriate as compared to agents who were acting with a neutral intent (see Figure 2). Similarly, agents who produced harmful outcome were condemned more harshly as compared to agents who did not. More importantly, MS patients did not differ from the control population in the degree to which they relied on outcome and intent information while deciding on how appropriate the agent's behavior was (see Figure 2).

[Figure 2 about here]

These main effects were qualified by an interaction effect between belief and outcome ($F(1,74) = 5.141, p = 0.026, p\eta^2 = 0.065, \omega^2 = 0.052$) but not a belief-by-outcome-by-group interaction ($F(1,74) = 0.057, p = 0.813, p\eta^2 = 0.001, \omega^2 \sim 0$). These results are indicative of

the fact that moral judgments about negative and neutral outcomes depended on whether these outcomes were produced with negative or neutral belief, and this dependence did not differ across groups (Figure 1). In particular, attempted harms were found to be less appropriate than the accidental harms based on the harmful intent ($ps < 0.001$). This oft-reported pattern of judgments from healthy population was also preserved in MS patients. One surprising result was that the behavior of accidental harm-doers was judged to be just as appropriate as the agents involved in neutral condition, but, more importantly, this was the case for both MS and HC groups (see Table 3).

There was also a main effect of group ($F(1,74) = 4.243, p = 0.043, p\eta^2 = 0.054, \omega^2 = 0.041$), such that MS patients made harsher judgments than the healthy controls (for additional details and scatter-plots, see Supplementary Text S5). Patients with MS generated higher average wrongness ratings than the HC group for 16 (out of 24) scenarios (see Supplementary Text S4). To investigate the source of this main effect, we carried out exploratory analyses in which we repeated the mixed effects ANOVA separately for each independent variable that differed between two groups, viz. EC, PD, DDF, DIF, and EOT (see Table 1). Only in the ANCOVA with EOT as the covariate did the main effect of group vanish ($\omega^2 = 0.036, p > 0.05$). In other words, although MS patients were not impaired in assessing the appropriateness of specific types of harmful acts (intentional, attempted, etc.), they were nonetheless harsher *in general* in judging third-party moral violations. Furthermore, once inter-group variance associated with externally-oriented cognition (EOT) was removed, the difference between MS patients and healthy controls was no longer significant.

3.5 Group differences in moral judgments about punishment

The pattern of results obtained for appropriateness judgments was also observed for punitive judgments (see Figure 3). A 2(belief) $\times$ 2(outcome) $\times$ 2(group) mixed ANOVA revealed a

main effect of belief ($F(1,74) = 714.63, p < 0.001, p\eta^2 = 0.906, \omega^2 = 0.906$), outcome ($F(1,74) = 22.495, p < 0.001, p\eta^2 = 0.233, \omega^2 = 0.233$), and also an interaction effect between belief and outcome ($F(1,74) = 11.196, p = 0.001, p\eta^2 = 0.131, \omega^2 = 0.131$). More importantly, none of these effects interacted with group factor (belief-by-group: $F(1,74) = 0.021, p = 0.886, p\eta^2 \sim 0, \omega^2 \sim 0$; outcome-by-group: $F(1,74) = 0.217, p = 0.643, p\eta^2 = 0.003, \omega^2 = 0.003$; belief-by-outcome-group: $F(1,74) = 0.205, p = 0.652, p\eta^2 = 0.003, \omega^2 = 0.003$).

[Figure 3 about here]

Put differently, while deciding on punishment for agents involved in moral situations, MS and HC groups relied to an equal extent on both belief and outcome information (for comparisons between different cases, see Table 3; for scatter-plots, see Supplementary Text S5).

Importantly, there was once again a main effect of group ($F(1,74) = 11.92, p < 0.001, p\eta^2 = 0.139, \omega^2 = 0.126$) such that MS delivered harsher moral judgments than controls (see Figure 3; for scatter-plots and additional details, see Supplementary Text S5). Strikingly, patients with MS generated higher average punishment ratings than the HC group for 22 (out of 24) scenarios (see Supplementary Text S4). In other words, not only did MS patients find the behavior of agents to be less appropriate than the HC group (as shown by the harsher appropriateness judgments across the board), but they were also more punitive than the control participants.

In exploratory analyses, we repeated this mixed ANOVA separately for each independent variable that differed between two groups, viz. EC, PD, DDF, DIF, and EOT (see Table 1), to probe source of this effect. All group differences remained significant in ANCOVAs with EC, PD, DDF, and DIF as covariates. Once again only in ANCOVA with EOT as the covariate, although the main effect of group was still significant ($p = 0.014$), none of the

planned group comparisons were significant (see Supplementary Text S5) and the effect size for the main effect of group was reduced by almost half (from $\omega^2 = 0.126$ to $\omega^2 = 0.063$). In other words, differences in EOT scores significantly accounted for elevated punitive judgments in MS patients than the controls.

3.6 Distinction between appropriateness and punishment judgments across groups

Prior work has shown that punishment judgments rely to a greater degree on outcome information as compared to appropriateness judgments (Cushman, 2008). In order to investigate whether this distinction was observed in the current study and whether it was preserved in the MS population, we conducted a 2(belief) $\times$ 2(outcome) $\times$ 2(type of judgment: appropriateness, punishment) $\times$ 2(group) mixed ANOVA. Indeed, this analysis revealed an outcome-by-judgment interaction ($F(1,74) = 4.623, p = 0.035, p\eta^2 = 0.059, \omega^2 = 0.083$), but no outcome-by-judgment-by-group interaction ($F(1,74) = 0.056, p = 0.813, p\eta^2 = 0.001, \omega^2 \sim 0$). To explore this effect further, we separately carried out a 2(outcome) $\times$ 2(judgment) $\times$ 2(group) mixed ANOVA for neutral (accidental harm vs. neutral cases) and negative (intentional vs. attempted harm cases) belief, but no outcome-by-judgment-by-group interaction was found either for neutral ($F(1,74) = 0.492, p = 0.485, p\eta^2 = 0.007, \omega^2 \sim 0$) or for negative belief ($F(1,74) = 0.077, p = 0.782, p\eta^2 = 0.001, \omega^2 \sim 0$).

To summarize, consistent with prior work (Cushman, 2008), participants assigned greater weight to information about the nature of outcomes (harmful or neutral) more heavily when deciding on punishment ($\omega^2 = 0.233$) as compared to judging appropriateness ($\omega^2 = 0.048$), irrespective of whether the intent was neutral or harmful. Interestingly, this distinction between moral judgments was preserved in the MS patients. The same conclusion can be

drawn from the percent of variance¹ explained by each factor for each type of judgment for the HC and MS groups (see Table 4).

[Table 4 about here]

This is a striking result as the MS patients scored lower than controls on EC and more on PD, and thus one would have expected them to be, respectively, either less or more reliant on outcomes while endorsing punishment judgments (vis-à-vis appropriateness judgments) as compared to controls. We note, however, that MS patients did not exhibit any reduction in the cognitive component of empathy (PT). This led us to suspect that this component of PT may play a crucial role in mediating the influence of harmful outcome on punishment judgments to a greater degree than appropriateness judgments. In other words, people may deliver harsher punishment versus appropriateness judgments for accidental harm-doers compared to neutral cases because of an empathic aversion stemming from victim perspective-taking (and not empathic concern or affective sharing) that motivates them to punish. This hypothesis predicts that, once the inter-individual variation in PT is accounted for, we should not observe interaction between the nature of outcome and type of moral judgment. Accordingly, we carried out a 2(belief) $\times$ 2(outcome) $\times$ 2(type of judgment) repeated measures ANOVA, yielding the expected interaction effect between outcome and type of judgment ($F(1,75) = 7.958, p = 0.006, p\eta^2 = 0.096$), but this interaction was no longer significant when PT was added as a covariate to the model ($F(1,74) = 0.130, p = 0.720, p\eta^2 = 0.002$).

¹Note that belief accounted for a surprising 84-91% of total variability, while outcome accounted for less than 1% of variance for both type of judgments. This contrasts with prior findings, which show that outcome explains up to 3% (for wrongness) or 20% (for punishment) of the total variation in moral judgment (e.g., Cushman, 2008). This pattern can also be discerned for appropriateness and punishment judgments where the agents who accidentally harmed someone were not judged more severely than the neutral cases by both groups. At present, we cannot distinguish whether this departure from past findings is due to the cultural settings in which the experiment was conducted (Argentina vs. US) or artefact of limited sample size ($n = 38$ vs. $n > 1000$) or differences in design (within-subject vs. between-subject).

Because this was an unexpected and novel result, we replicated this effect in a larger sample ($n = 113$) consisting of only healthy adults and with a larger battery of scenarios ($n = 36$). Replication of this effect was crucial for two additional reasons: (i) the moral judgments observed in the current sample were unusually dependent on intent information (as shown by equal punishment for accidental and neutral cases; cf. Table 4), and (ii) the scenarios were not counterbalanced across conditions. This new study addressed these concerns and successfully replicated the observed effects (see Appendix). Thus, we conclude that cognitive empathy (or PT) plays a key role in the relatively greater influence of outcome information on punishment versus appropriateness judgments.

3.7 Group differences in egocentric moral judgments

In order to assess group differences in egocentric moral judgments (also see Table 2), we ran a 2(belief)-by-2(outcome)-by-2(group) mixed ANOVA, which revealed a main effect of group such that MS patients reported that others would agree with their judgments to a greater degree than controls, irrespective of the type of scenario ($F(1,74) = 10.590, p = 0.002, \eta^2 = 0.054, \omega^2 = 0.112$). Thus, MS patients were more confident than the healthy controls that others would agree with their judgments (see Figure 4; for scatter-plots and additional details, see Supplementary Text S5). Indeed, patients with MS generated higher egocentric ratings than the HC group for 20 (out of 24) scenarios (see Supplementary Text S4).

[Figure 4 about here]

As with the punishment judgments, we carried out mixed ANCOVA separately for each covariate of interest (EC, PD, DDF, DIF, and EOT) to explore possible source of this effect. Once again, this analysis showed that the main effect of group was no longer significant ($\omega^2 \sim 0, p > 0.05$) only when EOT was added as a covariate to the model.

3.10 Correlation analysis: Relation between clinical variables, alexithymia, empathy, and moral judgments

Given that the sample size was limited for each group ($n = 38$), we provide preliminary analyses of the relationships among inter-individual differences in personality traits and clinical or dependent variables of interest in Supplementary Text S6.

3.11 Correlation analysis: Relation between appropriateness and punishment judgments

As would be expected, there was a high positive correlation between appropriateness and punishment judgments for a given condition, and these associations were not different between controls and MS patients, as shown by Fisher's Z-test (see Supplementary Text S7 for full analysis).

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3 **4. General Discussion**
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6 In the current study, we investigated how the known impairments in social cognition
7 processes in MS patients affect their moral judgments in contexts where information about
8 beliefs and outcomes conflict. The findings revealed a complex pattern of moral judgments
9 that deviate from healthy controls. We discuss each result individually in the following
10 subsections.
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18 **4.1 Preserved moral judgments in MS for specific acts**
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21 Past evidence has shown that both condemning agents who attempt but fail to cause harm and
22 exculpating agents who accidentally cause harm depend on recruitment of the ToM network
23 (e.g., Young et al., 2007; Young et al. 2010). This research has also underscored that mental
24 state reasoning is especially critical for the latter process – exculpation in the case of
25 accidents. In the case of a prepotent tendency to condemn the agent due to her causal
26 responsibility for the negative outcome (e.g., harm), a robust representation of the agent’s
27 innocent mental state is needed to override this response (Moran et al., 2011; Young et al.,
28 2007; Young & Saxe, 2009). This basic model also finds support from studies of clinical
29 populations with ToM impairments, like autism; individuals with autism properly assess the
30 wrongness of attempted harms but condemn accidental harms more (Buon et al., 2013b;
31 Moran et al., 2011; Salvano-Pardieu et al., 2015). Thus, we expected a similar pattern among
32 MS patients: elevated condemnation for accidental harm, but no difference for attempted
33 harm.
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51 In the current study, however, we did not observe any group difference for the accidental
52 harm cases. One possible explanation for this null result is that MS patients exhibit deficits
53 not only in ToM but also in empathic concern. Past research shows that more empathic
54 individuals (especially those scoring high on empathic concern) find accidental harms to be
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more wrong/unacceptable (Patil & Silani, 2014b; Trémolière & Djeriouat, 2016; Walter et al., 2012), whereas clinical populations with reduced empathy are more lenient in their moral judgments of accidental harms (alexithymia: Patil & Silani, 2014b; psychopathy: Young et al., 2012). Since there is likely to be no strong negative affect stemming from empathic aversion that would otherwise lead to increased condemnation of accidental harm cases in MS patients, following the two-process framework, no robust mental state representation is needed to forgive the agent. Specifically, the ToM and empathy deficits in MS patients likely exert mutually opposite influence on the final judgment that cancel each other out and leave their final appropriateness judgments relatively intact (reduced ToM → accidental harms more wrong; reduced empathic concern → accidental harms less wrong; final output → preserved moral assessment of accidental cases). In the case of MS, reduced empathy alongside ToM deficits therefore contribute to an apparently typical pattern of moral judgment. Note that of all the clinical populations studied so far on this task (autism, alexithymia, psychopathy), MS remains the only one which has deficits in ToM *and* empathy and not ToM *or* empathy. We also note that, although empathic reasoning and mentalizing are two independent processes that contribute their unique inputs (Cushman, 2008), these inputs need to be properly integrated for the final moral judgments (Buon et al., 2013a). Impairment in any one process (e.g. mentalizing) will leave the input from other process intact (e.g. although mentalizing is impaired in autism, the input from empathizing is preserved), and this will affect the final output (increased condemnation for accidents). We propose that, in case of MS (as compared to HC), since both processes are impaired, the final output is comparable to that in HC.

4.2 Role of EOT in harsher global moral judgments in MS

Although MS patients did not differ from controls in terms of their reliance on belief and outcome information while endorsing moral judgments, they nevertheless found the third-

party behaviors (regardless of the type of scenario) to be less appropriate and more deserving of punishment and this effect was driven by EOT. Note that neither ToM nor empathy deficits predict this pattern for reasons mentioned in the preceding section (Section 4.1).

One possible explanation is that this harsher moral attitude towards others' behavior stems from the affective disorders (depression, anxiety, and fatigue) that are prevalent and persistent in MS patients. (e.g., Bodini et al., 2008; Chahraoui et al., 2014; Gay et al., 2010; for a review see Sá, 2008). MS is a chronic disease without a cure that usually strikes in the most formative years of one's life (20-40 years). Additionally, the long-term course of MS highly varies from patient-to-patient; some patients can go without an attack for years, while others can go from being perfectly healthy to having a rapid physical and cognitive decline in a short period of time. This unpredictable and ambiguous nature of progression in MS is considered to be the fountainhead of affective disorders in MS (Chahraoui et al., 2014). When patients are in remission phase, they tend to be constantly in a state of fear and anxiety due to the highly unpredictable nature of their disease course, uncertain of when the next attack (relapse) can flare up; in addition, during the course of a relapse, they worry about possible imminent disability. Extant research (Chahraoui et al., 2014) suggests that one possible defensive strategy that patients use to cope with the traumatic experience from past attacks (Sá, 2008) is to develop a trait-like pattern of orienting their thoughts on the external events rather than introspecting on their inner experiences and feelings (i.e. EOT). Although alexithymia, in general, has been argued to be an inhibition mechanism that acts as an anti-stress fortification both in healthy and clinical populations, with different dimensions of alexithymia playing a role in specific situations or in response to specific stimuli (Davydov et al., 2010), EOT seems to be the locus of this mechanism. Studies using multiple approaches have revealed evidence for the role of EOT as an adaptive strategy that leads to the avoidance of self-reflection and thereby blunts the influence of negative arousal states to protect

individual in the long term from stress (for a more detailed discussion, see Demers & Koven, 2015). For example, EOT has been found to distract attention from self-oriented ruminations towards external environment, thereby disrupting repetitive negative thoughts that exacerbate sad or depressed mood while protecting the individual from negative arousal states (Luminet et al., 2004). EOT is also associated with an inability to adopt an experiential mode of thinking (Di Schiena et al., 2011). Psychophysiological studies demonstrate the moderating role of EOT in decoupling subjective feeling states from physiological arousal states (Davydov et al., 2013), and behavioral studies demonstrate that DDF and DIF, but not EOT, predict levels of depression and anxiety in healthy controls (Hendryx et al., 1991). Additional evidence supporting role of EOT as a cognitive, adaptive strategy comes from clinical populations with affective disorders in which DDF and DIF - but not EOT - account for variance in depression and anxiety, e.g., posttraumatic stress disorder (Monson et al., 2004), childhood abuse and neglect (Güleç et al., 2013), borderline personality disorder (Evren et al., 2012), depressive and anxiety disorders (Marchesi et al., 2000; Saarijärvi et al., 2001), etc. Similarly, implication of EOT as an adaptive strategy in MS population explains why DIF and DDF factors of alexithymia, but not EOT, are found to predict depression, anxiety, and fatigue in MS patients (Bodini et al., 2008; Chahraoui et al., 2014).

Given this reduced ability to effectively process their own emotional states and a cognitive style that is externally oriented, MS patients are expected to have poor emotion differentiation skills (Lindquist & Barrett, 2008), which would lead them to represent their affective states only in global terms. Therefore, they would not be able to disentangle affect *integral* to the scenario (in response to harmful intent and/or outcome) from *incidental* affect (cf. Cameron, Payne, & Doris, 2013). Both integral and incidental affect have been well-known to influence judgment and decision making (for a review, see Västfjäll et al., 2016). Thus, the prevalent negative mood states in MS can be misattributed by patients to the

experimental stimulus and this negative emotional arousal, in turn, would amp up the severity of moral condemnation (Cheng, Ottati, & Price, 2013). Västfjäll and colleagues have argued that the incidental affect can have especially greater bearing on judgments when the incidental affect is salient but the source awareness is weak. This is indeed the case for MS patients who experience salient negative affective states, but the EOT strategy reduces their awareness of source of this affect.

The current data support this hypothesis given that the main effect for the appropriateness judgments was not significant and significantly reduced for punishment judgment only after controlling for the group differences in the EOT component of alexithymia (Bodini et al., 2008; Chahraoui et al., 2014). Thus, the strategy that MS patients employ to protect themselves from distressing affective states seems to also prevent them from insulating their moral judgments² from the prevalent, incidental negative affect that leads them to make harsher judgments as compared to healthy controls (see Figure 5 for schematic illustration). In other words, the harsher moral attitudes in MS seem to be not due to MS affectation per se, but due to comorbid alexithymia, especially its subcomponent EOT.

[Figure 5 about here]

Note that this interpretation still does not explain why the same effect was relatively weak for wrongness judgments, i.e. why the incidental negative arousal affected punishment judgments to a greater extent (effect size: $\omega^2 = 0.126$) than wrongness judgments (effect size: $\omega^2 = 0.041$). We propose that this is due to the fact that blame/punishment judgments in

²The current results motivated us to reanalyze findings from a previous study (Gleichgerricht et al., 2015), which showed that MS patients were less utilitarian on moral dilemmas than healthy controls due to increased emotional reactivity. This analysis showed that the increased emotional reactivity in MS was no longer significantly different than controls once variance associated with EOT (and not DDF or DIF) was controlled for. Furthermore, higher scores on EOT were predictive of more egocentric moral judgments. These results further bolster our claim that the aberrant moral judgments in MS stem from heightened externally oriented cognitive style that prevents them from properly harnessing affective responses for making moral judgments.

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3 general rely to a greater degree on affective processes than wrongness/appropriateness
4 judgments, as shown by both neuroimaging (Buckholtz et al., 2008, 2015; Treadway et al.,
5 2014) and behavioral (Cushman, 2008; also see Malle, Guglielmo, & Monroe, 2014) studies.
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10 **4.3 Increased egocentric moral attitudes in MS**

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12 Interestingly, MS patients were more likely to report that other people would evaluate third-
13 party moral violations the same way as they do than the healthy controls. In other words, MS
14 patients displayed more self-oriented moral attitudes and could not foresee that others may
15 hold dissenting moral opinions about the behavior of agents in these scenarios. Interestingly,
16 this effect was also not significant once group differences in EOT were accounted for. One
17 possible explanation for this effect comes from the observation that higher EOT scores
18 indicate reduced ability to engage in symbolic thought (Herbert et al., 2011), which has been
19 argued to be important for appreciating others' perspectives and even necessary for acquiring
20 the ability to engage in mentalizing (Tuch, 2011). Thus, the increased tendency to engage in
21 externally-oriented thinking coupled with the observed difficulties in MS to properly
22 represent others' mental states may lead MS patients to hold a more egocentric moral frame-
23 of-reference and assume that others share their moral views. The same effect has also been
24 observed for moral judgments on moral dilemmas where MS patients exhibited increased
25 egocentric moral judgments (Gleichgerricht et al., 2015).
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45 **4.4 Cognitive empathy and the influence of moral luck on punishment vs. wrongness**

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47 Although there is evidence to support the claim that outcomes influence punishment
48 judgments to a greater extent than wrongness/appropriateness judgments (Cushman, 2008;
49 Cushman et al., 2009; Cushman et al., 2013), the exact psychological mechanism that
50 underpins this process remains unexplored. Neurobiological models of punishment posit
51 empathic aversion stemming from victim distress to be a source of aversive reinforcement
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that motivates punishment with the ultimate goal of changing the agent’s behavior (Martin & Cushman, 2016; Seymour, Singer, & Dolan, 2007). Nonetheless, so far, no further distinction has been made as to precisely which component of empathy (cognitive, affective, or motivational) is involved in conveying the influence of harmful outcomes on punishment judgments.

In the current study, MS patients exhibited reduced motivational empathy (i.e., EC) and increased affective sharing (i.e., PD) as compared to healthy controls, and yet they did not differ from the control population in terms of punishment/wrongness distinction for moral luck contrast cases. In other words, both MS patients and healthy controls increased punishment for accidents (versus neutral cases) more than they increased wrongness judgments. This led us to hypothesize that the cognitive component of empathy (PT), intact in MS patients, may subserve the differential influence of outcomes on punishment versus wrongness judgments (see Figure 5). Accordingly, in the current sample, once the inter-individual variation in PT was accounted for, we no longer observed the punishment/wrongness distinction for the key cases. We additionally replicated this effect in a large independent sample and with an expanded set of stimuli in a healthy population from a different cultural setting (see Appendix). This study also showed that removing the variance associated with only PT, and not EC or PD, led to disappearance of interaction between nature of outcome (harmful or not) and type of moral judgment (punishment or wrongness).

Thus, based on the current findings, we propose a refinement of the existing two-process model for intent-based judgments by showing that cognitive empathy or victim perspective-taking is the precise source of affect that motivates a third-party judge to punish agents for causing harmful outcomes. We illustrate the proposed modifications to the existing two-process model in Figure 5.

4.5 Implications

The clinical meaningfulness of the current findings requires further exploration, but we argue here that this peculiar pattern of moral judgments can have implications for proper social functioning in MS. Note that not only did MS patients judge others' behavior to be more wrong (or less appropriate) and were more punitive than controls, but they were also more confident that others would evaluate moral situations the same as they do. Although we have provided evidence here only for elevated third-party punishment judgments, we would also expect the same pattern to hold for second-party punishments, more common in day-to-day life, as these judgments recruit similar underlying neural processes (Buckholtz et al., 2008; Yu et al., 2015). This moral profile increases the risk for social dysfunction in MS patients as the individuals interacting with them (their social support system, for instance) may be put off by their punitive attitudes (especially in neutral cases where agents neither intend nor cause harm). In turn, such interpersonal conflict may be exacerbated due to MS patients' apparent inability to appreciate others' moral views during such moral disagreements. This dynamic may lead to greater social exclusion in MS patients, with negative consequences for psychological and social well-being (Montel & Bungener, 2007; Phillips et al., 2009). Thus, intervention strategies designed to improve quality of life in MS patients by increasing their social participation should focus on bringing patients' exaggerated punitiveness to their notice and also encouraging them to see moral situations from others' point-of-view during moral disagreements. This may be achieved by targeting specific socioaffective and sociocognitive processes, such as working on attenuating tendencies to engage in EOT to regulate negative affective states and adopt more adaptive strategies like voluntary attention to emotions (Boden & Thomson, 2015). Additionally, it is known that being aware of the source of the negative affective states can limit its influence on judgments (Västfjäll et al., 2016). Thus, intervention focused on increasing MS patients' awareness of their affective

states can make them distrust using this affect as information and reduce severity of their moral condemnation. This prediction stemming from the current study can be investigated in future studies.

4.6 Limitations

4.6.1 Methodological

The conclusions drawn from this study need to be interpreted with the following limitations in mind. First, as noted in the Methods section, the scenarios were not randomized across conditions, and thus we could not assess validity of the results from the subjects-level analysis of moral judgments in an item-level analysis to ascertain that main results were not being driven by few artifactual items (cf. McGuire et al., 2009). Second, we relied heavily on the IRI questionnaire to assess different components of empathy, but the IRI has been argued to be a measure with limited ecological validity (Ickes, 2009), and the robustness of the current results needs to be demonstrated with other available instruments that also tap into these different components (e.g. QCAE: Questionnaire of Cognitive and Affective Empathy; Reniers et al., 2011). Third, the current study did not directly assess ToM deficits in MS population, but we do note the substantial literature demonstrating these deficits in MS patients (Banati et al., 2010; Charvet et al., 2014; Henry et al., 2009, 2011; Kraemer et al., 2013; Ouellet et al., 2010; Pöttgen et al., 2013; Roca et al., 2014). Fourth, this was an exploratory study that found EOT to play an important role in moral cognition, and thus future studies need to further explore the exact mechanism by which this dimension of alexithymia interacts with incidental affect and how this interaction has downstream consequences for moral judgments. Fifth, the Spanish version of the TAS-20 used in the current study was validated in the general population (Martínez Sánchez, 1996) but not in the MS population. One recent study (Fernández-Jiménez et al., 2013) assessed the factor

structure of TAS in MS population, and future studies should use this version to study cognition in MS. Lastly, it is important to keep in mind that these results are statistical in nature and are not the case reports of profound social dysfunction and moral deficits associated with brain pathology in the literature. This limitation is also important in the light of work that reveals how people usually exhibit discrepancy between judgments they provide on hypothetical scenarios and their behavior in a more ecologically valid setting (Patil et al., 2014). Hence, the translation of the findings will need clear correlation to real world functioning parameters.

4.6.2 Statistical

Note that our main conclusions are based on ANCOVAs with covariates that differed between groups (e.g., EOT; in Sections 3.4, 3.5, and 3.7). This has been argued to be a weak procedure as there is no meaningful way to adjust for covariates when they share significant variance with the group factor; in such cases, removing the variance associated with the covariate also alters the grouping variable substantively (Miller & Chapman, 2001). But note that in the current study, these ANCOVAs were carried out exclusively to *explore* the source of the observed effect and not as *hypothesis testing*, a specific case where this practice has been considered to be defensible (Huitema, 2011, p. 212). We did not *a priori* predict either increased punitiveness or moral egocentrism, nor did we have any prediction for a possible role of EOT in explaining this effect; ANCOVA was carried out to explore the possible role of different independent variables with clinical relevance to the MS population. The current findings should be confirmed in a future hypothesis-driven (rather than exploratory) study where ANCOVA can be circumvented by matching the MS group for EOT scores with healthy controls. In such study, the current proposal would predict that the MS patients will not exhibit harsher moral condemnation profile.

5. Conclusions

In the current study, we investigated intent-based moral judgments in MS patients to assess how the observed deficits in social cognition in this population affect their moral evaluations. We found that, although MS patients judged specific types of third-party moral violations appropriately by integrating information about beliefs and outcomes, they nonetheless judged behaviors to be *generally* more wrong, they were more punitive, and they were more confident about their judgments as compared to healthy controls. We argued building on past work that this increased moral condemnation in MS is likely to be due to unrestrained influence of prevalent, *incidental* negative affective states whose influence on moral judgments is not differentiated from *integral* affect stemming from stimuli due to externally oriented cognitive style. Moreover, more egocentric moral attitudes in MS patients may also arise from confluent effect of self-centered perspective due to external preoccupations and reduced perspective-taking due to ToM deficits. The distinction between wrongness and punishment judgments for moral luck cases was found to be preserved in MS patients due to an intact cognitive capacity for victim perspective-taking. Finally, we point out the translational value of the current research for therapeutic practices designed to improve quality of life in MS patients.

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Table 1. *Group differences in levels of alexithymia and empathy.* Values in parentheses denote standard deviation

Variable	MS (<i>n</i> = 38)	HC (<i>n</i> = 38)	<i>t</i>	<i>df</i>	<i>p</i>	Cohen's <i>d</i>	BF10
DDF	13.95 (4.17)	10.89 (3.73)	3.364	73.13	0.001	0.772	25.77
DIF	15.97 (7.57)	11.45 (4.67)	3.137	61.62	0.003	0.72	14.34
EOT	27.45 (4.83)	15.82 (4.23)	11.175	72.75	< .001	2.564	2.389e +14
TAS-20	57.37 (13.40)	38.16 (9.71)	7.156	67.46	< .001	1.642	1.563e +7
PT	3.602 (0.642)	3.425 (0.645)	1.196	74	0.235	0.274	0.44
F	2.891 (0.795)	3.41 (0.846)	-2.754	73.72	0.007	-0.632	5.762
EC	3.342 (0.67)	3.816 (0.624)	-3.188	73.63	0.002	-0.731	16.287
PD	3.066 (0.724)	2.544 (0.812)	2.957	73.07	0.004	0.678	9.233

Table 2. Mean (SD) for appropriateness, punishment, and egocentric moral judgments for each group.

Type of judgment	Scenario	Group	Mean (SD)
Appropriateness	neutral	HC	2.259 (1.029)
		MS	2.618 (0.918)
	accidental	HC	2.197 (0.856)
		MS	2.693 (1.125)
	attempted	HC	6.224 (0.735)
		MS	6.289 (1.037)
Punishment	intentionl	HC	6.509 (0.517)
		MS	6.645 (0.653)
	neutral	HC	1.544 (0.629)
		MS	2.105 (0.974)
	accidental	HC	1.684 (0.75)
		MS	2.259 (0.992)
	attempted	HC	5.004 (1.271)
		MS	5.548 (1.319)
	intentionl	HC	5.509 (1.167)
		MS	6.18 (1.055)
Egocentric	neutral	HC	79.123 (16.184)
		MS	80.246 (13.561)
	accidental	HC	71.382 (15.582)
		MS	79.803 (13.961)
	attempted	HC	67.75 (14.766)
		MS	82.522 (14.209)
	intentionl	HC	76 (13.486)
		MS	87.93 (13.286)

Table 3. Bonferroni-corrected post-hoc comparisons between neutral/negative belief and neutral/negative outcome cases for MS and HC groups.

Comparison	Group	Appropriateness		Punishment	
		Mean difference	<i>p</i> -value	Mean difference	<i>p</i> -value
accidental-neutral	HC	-0.061	0.681	0.14	0.436
	MS	0.075	0.667	0.154	0.484
intentional-attempted	HC	0.285	0.042	0.504	< 0.001
	MS	0.355	0.020	0.632	0.002
attempted-neutral	HC	3.965	< 0.001	3.461	< 0.001
	MS	3.671	< 0.001	3.443	< 0.001
intentional-accidental	HC	4.311	< 0.001	3.825	< 0.001
	MS	3.952	< 0.001	3.921	< 0.001

Table 4. *Percentage of total variability explained by belief and outcome factors and their interactions for each type of moral judgment and for each group.*

Factor	appropriateness		punishment	
	HC	MS	HC	MS
belief	91.68	87.03	87.07	84.00
outcome	0.07	0.28	0.68	0.95
belief-by-outcome	0.16	0.12	0.22	0.35
Error	8.09	12.58	12.03	14.69

Figure Captions

Figure 1. Four types of possible harms (conditions) from a 2 (belief: neutral, harmful) × 2 (outcome: neutral, harmful) design. One example is also shown in the form of an abbreviated story. Note that in the current study each scenario appeared only in one condition. See Supplementary Text S2 to access full list of stories and the conditions in which they appeared.

Figure 2. Moral judgments about appropriateness of behavior of moral agents by MS patients and healthy controls on a 7-point Likert scale (1: *completely appropriate*, 7: *completely inappropriate*) for different types of harms: neutral case (neutral belief, neutral outcome), accidental harm (neutral belief, negative outcome), attempted harm (negative belief, neutral outcome), and intentional harm (negative belief, negative outcome). Error bars represent 95% confidence intervals.

Figure 3. Punitive judgments endorsed by MS patients and healthy controls on a 7-point Likert scale (1: *no punishment*, 7: *severe punishment*) for different types of harms: neutral case (neutral belief, neutral outcome), accidental harm (neutral belief, negative outcome), attempted harm (negative belief, neutral outcome), and intentional harm (negative belief, negative outcome). Error bars represent 95% confidence intervals.

Figure 4. Egocentric moral judgments by MS patients and healthy controls to assess the expected degree of congruency between one’s own response and others’ responses ("Out of 100 people answering to this scenario, how many do you think would answer like you?"), for different types of harms: neutral case (neutral belief, neutral outcome), accidental harm (neutral belief, negative outcome), attempted harm (negative belief, neutral outcome), and intentional harm (negative belief, negative outcome). The responses were collected using a

continuous scale whereby participants could choose any number from 0 to 100. Error bars represent 95% confidence intervals.

Figure 5. Schematic depiction of a two process model for intent-based moral judgement and the extensions we have proposed based on results from the current study. A causal reasoning process determines the causal role of the agent only in the presence of harmful consequence and condemns the agent, but remains silent in absence of harmful outcome. A mental state reasoning process probes for culpable mental states and condemns the agent when such state is found. Both of these processes are source of negative affect *integral* to the moral violation and provide independent evaluations that are integrated to a different degree for different categories of judgments: punishment judgments rely to a greater degree on information about outcomes than appropriates judgments (represented by solid and dotted lines) and the current study proposes that this influence is mediated by cognitive empathy. In other words, outcomes matter more for punitive attitudes only to the degree that people engage in representing victim suffering. Furthermore, based on current findings, we propose that the *incidental* negative affect (e.g., stress, depression, etc.) can amplify the severity of the final integrated judgment in case of poor emotional differentiation skills, stemming from externally focused cognition (like in MS), that insulate judgment from incidental affect and ensure reliance on only the affect integral to the situation. Adapted from Cushman (2008).

Figure 6. Results from the replication study. (a) Proportion of within-condition variability explained by each factor for the acceptability judgment. (b) Proportion of within-condition variability explained by each factor for the punishment judgment. (c) Direct comparison of mean moral judgment for each condition grouped by type of judgment.

Appendix: Replication study

Participants:

A total of 113 healthy native Italian-speaking community members (70 females) came to the lab to participate in this study and were financially compensated for their time. All participants provided written informed consent. Average age was 24.42 (SD = 5.55, median = 23) years, with a range of 18 to 57.

Experimental stimuli:

Experimental stimuli consisted of four different sets or versions of 36 unique vignettes for a total of 144 stories and participants were assigned different versions in Latin-square arrangement. Each participant saw one variation of each scenario, for a total of 36 stories. All scenarios were adapted in Italian from Young et al. (2010). The four versions were the result of a 2-by-2 design where the factors *belief* (neutral, negative) and *outcome* (neutral, negative) were independently varied such that agents in the scenario produced either a neutral outcome or a harmful outcome while acting with the belief that they were causing either a neutral outcome or a harmful outcome (see Figure 6).

[Figure 6 about here]

Each scenario lasted for 32 s and consisted of four cumulative segments (each lasting for 8 s): (i) *background*: this stem was common to all variations and provided settings in which the story took place; (ii) *foreshadow*: this segment foreshadowed whether the outcome will be neutral or harmful; (iii) *belief*: this segment provided information about whether the agent was acting with a neutral or harmful belief; (iv) *consequence*: this final segment revealed outcome of the agent’s action. All the story text was then removed and replaced with the question and response scale.

After reading each scenario, participants provided two types of moral judgments which were presented in a randomized order:

(i) *acceptability*: “How morally acceptable was [the agent]’s behavior?” (from “Not at all acceptable” to “Completely acceptable”);

(ii) *punishment*: “How much punishment does [the agent] deserve for his/her behavior?” (from “No punishment” to “The most severe punishment”).

After each story was presented, participants responded using computerized visual analog scales (VAS), implemented as horizontal onscreen bar and responses were later converted to standardized scores with [min, max] of [0, 20]. The acceptability scores were reverse-scored so that higher score for the two questions indicated less acceptable behavior and more punishment. Participants had 6 s to respond to each question.

Participants then completed the Italian-validated version of the IRI (Albiero et al., 2006).

Note that all subscales of Italian version had the same items as that in English version.

Data analysis:

All effect sizes have been reported based on recommendations in Lakens (2013).

Results:

For the acceptability judgments, the belief and outcome factors respectively accounted for 65% and 18% of the variability (Figure 6(a)), while for the punishment judgments these factors accounted for 44% and 37% of the variance (Figure 6(b)). Thus, as predicted by the two-process model, punishment judgments were dependent to an equal degree on both belief and outcome information, while the acceptability judgments depended to a large extent on the mental state information.

ANOVA analysis also revealed the same pattern: there was a main effect of belief and outcome for both acceptability (belief: $F(1,112) = 835.08, p < 0.001, p\eta^2 = 0.882$, 90% CI [0.849, 0.902], $\eta_G^2 = 0.377$; outcome: $F(1,112) = 448.63, p < 0.001, p\eta^2 = 0.800$, 90% CI [0.746, 0.835], $\eta_G^2 = 0.063$) and punishment (belief: $F(1,112) = 516.77, p < 0.001, p\eta^2 = 0.822$, 90% CI [0.773, 0.853], $\eta_G^2 = 0.377$; outcome: $F(1,112) = 714.43, p < 0.001, p\eta^2 = 0.864$, 90% CI [0.827, 0.888], $\eta_G^2 = 0.063$) judgments such that behavior of agents who produced harmful outcome or who harbored intention to harm was condemned and punished more severely (see Figure 6(c)). Additionally, there was also a belief-by-outcome interaction for acceptability ($F(1,112) = 62.80, p < 0.001, p\eta^2 = 0.359$, 90% CI [0.243, 0.456]) but not for punishment ($F(1,112) = 2.71, p = 0.102$) judgments.

More important to the current replication study, we observed a outcome-by- type of judgment interaction in a 2 (belief) $\times$ 2 (outcome) $\times$ 2 (type of judgment) repeated measures ANOVA ($F(1,112) = 81.893, p < 0.001, p\eta^2 = 0.422$). This interaction signifies that participants relied on outcome information to a different degree while evaluating acceptability of and punishment for the behavior under consideration. In particular, punishment for behavior of an agent with harmful intent was *reduced* significantly more than acceptability when she failed as compared to when she succeeded in producing harmful outcome ($F(1,112) = 88.29, p < 0.001, p\eta^2 = 0.441$, 90% CI [0.327, 0.530]). Similarly, severity of endorsed punishment for an agent who accidentally produced harm while acting under false belief was *increased* significantly more than acceptability as compared to a neutral case ($F(1,112) = 10.28, p = 0.002, p\eta^2 = 0.084$, 90% CI [0.020, 0.173]).

To address the primary replication study hypothesis, we re-ran the 2 (belief) $\times$ 2 (outcome) $\times$ 2 (type of judgment) repeated measures ANOVA but with PT as a covariate. This model did not reveal any significant interaction between outcome and judgment type ($F(1,111) = 2.393$,

$p = 0.125$, $p\eta^2 = 0.021$). But this interaction remained significant if EC ($F(1,111) = 6.016$, $p = 0.016$, $p\eta^2 = 0.051$) or PD ($F(1,111) = 7.355$, $p = 0.008$, $p\eta^2 = 0.064$) was added as a covariate to the model.

Thus, we replicated our finding from the main study that once the inter-individual differences in cognitive empathy are accounted for, outcomes no longer have more influence on punishment judgments than the wrongness judgments. In other words, perspective-taking, and not empathic concern or personal distress (which are motivational and affective aspects of empathy, respectively), seems to be at the core of why moral luck has a greater bearing on punishment judgments than wrongness judgments. This also explains why MS patients showed the usual wrongness/punishment distinction in terms of reliance on outcome information; although MS patients showed reduced EC and increased PD, they did not differ from the control population in their capacity to engage in cognitive empathizing.

		Intent	
		neutral	negative
Outcome	neutral	neutral case: X puts peanuts in Y's food thinking Y is not allergic to peanuts; Y eats and enjoys his meal	attempted harm: X puts peanuts in Y's food, even though X is aware that Y is allergic to peanuts; Y eats and enjoys his meal
	negative	accidental harm: X puts peanuts in Y's food thinking Y is not allergic to peanuts; Y suffers anaphylactic shock	intentional harm: X puts peanuts in Y's food, even though X is aware that Y is allergic to peanuts; Y suffers anaphylactic shock

Review Only

Figure 2
211x158mm (96 x 96 DPI)

Figure 3
211x158mm (96 x 96 DPI)

Figure 4
211x158mm (96 x 96 DPI)

Figure 5

395x295mm (72 x 72 DPI)

Supplementary information for “Elevated moral condemnation of third-party violations in multiple sclerosis patients”

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Supplementary Text S1: Empathy and Alexithymia measurement

Empathy

The Interpersonal Reactivity Inventory (IRI) (Davis, 1983) is a widely used self-report questionnaire to assess specific aspects of dispositional empathy. Participants reported agreement with statements on a 5-point Likert scale (1: *never true for me*, 5: *always true for me*). The four subscales consisted of: (1) fantasy (F) scale, which measures the propensity to identify with fictional characters (e.g., “I daydream and fantasize, with some regularity, about things that might happen to me.”); (2) perspective taking (PT) scale, which indexes the dispositional tendency to consider others’ perspective in interpersonal interactions (e.g., “I sometimes find it difficult to see things from the “other guy’s” point of view.”; reverse-scored); (3) empathic concern (EC) scale, which measures the *other-oriented* tendency to experience feelings of warmth, compassion, and concern for other people (e.g., “I often have tender, concerned feelings for people less fortunate than me.”); (4) personal distress (PD) scale, which measures the *self-oriented* tendency to feel personal unease and discomfort in reaction to the emotions of others (e.g., “Being in a tense emotional situation scares me.”). We used the version of IRI validated for Spanish population (Pérez-Albéniz et al., 2003). One noteworthy departure in this version of IRI from the standard one is that the EC scale has 8 items, while the PD scale has 6 items because one item (“When I see someone get hurt, I tend to remain calm.”; reverse-scored) from original PD scale was found to load on EC scale in the validation study. As different subscales had different number of items, we used mean (instead of sum) of items as a measure of empathy.

Alexithymia

To assess levels of trait alexithymia, we used the validated Spanish version of the Toronto Alexithymia Scale-20 (TAS-20) questionnaire (Bagby, Taylor, & Parker, 1994; Spanish version: Martínez Sánchez, 1996), which has been argued to be the best current measure overall for assessing alexithymia due to its sound reliability, validity, and broad

generalizability (Timoney & Holder, 2013). The TAS-20 is a 20-item scale that consists of three subscales: Difficulty Describing Feelings (DDF, 5 items, e.g., “I am often puzzled by sensations in my body.”), Difficulty Identifying Feelings (DIF, 7 items, e.g., “I often don’t know why I am angry.”), and Externally-Oriented Thinking (EOT, 8 items, e.g., “I can feel close to someone, even in moments of silence.”; reverse-scored). Items were rated using a 5-point Likert scale (1: *strongly disagree*, 5: *strongly agree*).

Internal reliability

Cronbach’s alpha for subscales of TAS and IRI questionnaires split by group. (For notes on abbreviations, see Table 1 from main text). Group differences in Cronbach's alphas were investigated using `cocron` package in R (<http://comparingcronbachalphas.org/>) which implements inferential statistics on alphas.

Variable	Cronbach's alpha		Item count	$\chi^2(1)$	<i>p</i>
	HC	MS			
DDF	0.637	0.402	5	1.4986	0.2209
DIF	0.781	0.848	7	0.9113	0.3398
EOT	0.495	0.536	8	0.0512	0.8210
TAS-20	0.783	0.819	20	0.2742	0.6005
EC	0.805	0.531	8	5.3088	0.0212
PD	0.797	0.589	6	3.1787	0.0746
PT	0.725	0.541	7	1.7825	0.1818
F	0.843	0.615	7	5.3328	0.0209

As can be seen from the above table, although internal reliability was consistently lower for the MS group as compared to the HC group, it was within acceptable range. Internal reliability differed significantly between groups only for fantasy and empathic concern scale, but not on PT which was more crucial for the effects observed.

The only unacceptable value for alpha was observed for DDF subscale of TAS for the MS group. We ensured that the significant difference between groups was not due to poor internal reliability by re-computing DDF after dropping the item loading poorly (“I am able to describe my feelings easily”; reverse-scored) on this subscale and redoing the test. This

reanalysis upheld the original result: MS patients showed higher levels of DDF than HC group (MS = 9.97, HC = 8.05, $t(63.16) = 2.207$, $p = 0.031$, $BF_{10} = 1.885$).

Supplementary Text S2

The scenario-by-condition breakdown. The exact wording of the English version of scenarios can be found in Young et al. (2010) using the corresponding label. The Spanish translations of the scenarios are available from authors on request.

Scenario no.	Name	Condition
SC1	Popcorn	intentional
SC2	River	attempted
SC3	Coffee	neutral
SC4	Logan Airport	attempted
SC5	Vitamin	accidental
SC6	Safety town	intentional
SC7	Harness	attempted
SC8	Laboratory	neutral
SC9	Sesame	accidental
SC10	Meatloaf	attempted
SC11	Jellyfish	neutral
SC12	Veternerian	intentional
SC13	Latex	accidental
SC14	Mushrooms	neutral
SC15	Motorboat	accidental
SC16	Tracks	attempted
SC17	Igloo	intentional
SC18	Bike	accidental
SC19	Ham	intentional
SC20	Pool	neutral
SC21	Asthma	accidental
SC22	Chairlift	intentional
SC23	Iron	attempted
SC24	Bridge	neutral

Supplementary Text S3

Internal reliability for each scenario condition and for each type of moral judgment.

MS group and healthy controls showed poor internal reliability for appropriateness judgments on neutral and intentional harm conditions, respectively, and results from these conditions

should be interpreted while keeping in mind this limitation. All other cells showed satisfactory internal reliability.

Judgment	Group	neutral	accidental	attempted	intentional
Appropriateness	HC	0.678	0.612	0.682	0.398
	MS	0.395	0.549	0.810	0.516
Punishment	HC	0.575	0.624	0.860	0.849
	MS	0.623	0.553	0.779	0.727
Egocentric	HC	0.807	0.836	0.798	0.767
	MS	0.783	0.811	0.848	0.861

Note that it can be argued that the lack of difference between appropriateness of neutral and accidental scenarios was due to the low internal reliability for appropriateness judgments (see Table 3 in the main text). This explanation seems unlikely since the same pattern was observed not only for appropriateness judgment but also for punishment judgments, which were associated with high internal reliability. We therefore maintain that the result for appropriateness judgments may reflect genuine overreliance on intent information in the current sample, which led to especially lenient judgments of accidental harms.

Supplementary Text S4

Average moral judgments about appropriateness, punishment, and moral relativism for MS and HC groups for individual items belonging to neutral, accidental, attempted, and intentional harm conditions.

Scenario	Group	Appropriateness	Punishment	Egocentric
SC1	HC	6.632	4.842	68.58
	MS	6.605	5.974	79.66
SC2	HC	6	4.079	56.92
	MS	6.053	5.289	76.68
SC3	HC	2.368	1.526	77.29
	MS	2.263	1.842	83.24
SC4	HC	6.421	5.658	67.79
	MS	6.447	5.895	85.5
SC5	HC	2.263	1.368	65.89
	MS	3.158	2.553	79.92
SC6	HC	6.921	5.789	79.58
	MS	6.763	6.368	90.97
SC7	HC	6.5	5.053	68.71
	MS	6.263	5.842	86.42
SC8	HC	1.868	1.474	84.92
	MS	1.974	1.921	87.08
SC9	HC	1.711	1.684	77.97
	MS	1.816	1.526	86.97
SC10	HC	6.5	5.447	71.95
	MS	6.553	5.842	87.5
SC11	HC	1.868	1.368	82.13
	MS	2.632	1.974	81.18
SC12	HC	6.974	6.579	88.47
	MS	6.789	6.368	92.21
SC13	HC	1.368	1.263	81.29
	MS	1.816	1.579	90.39
SC14	HC	3.053	2.079	71.24
	MS	3.842	3.053	73.11
SC15	HC	2.395	1.737	76.5
	MS	3.579	2.947	72.37
SC16	HC	6.5	5.711	73.76
	MS	6.447	5.921	82.61
SC17	HC	5.763	4.868	68.53
	MS	6.368	5.684	84.47
SC18	HC	2.237	2	66.89
	MS	2.184	2.053	77.71
SC19	HC	6.026	4.763	70.03
	MS	6.658	5.789	87.74
SC20	HC	2.289	1.447	77.63
	MS	2.605	1.947	77.16
SC21	HC	3.211	2.053	59.74
	MS	3.605	2.895	71.45
SC22	HC	6.737	6.211	80.82
	MS	6.684	6.895	92.53
SC23	HC	5.421	4.079	67.37
	MS	5.974	4.5	76.42
SC24	HC	2.105	1.368	81.53
	MS	2.395	1.895	79.71

Supplementary Text S5

In this section, we provide:

(1) Scatter-plots for moral judgments for MS and HC groups. Higher score on the Y-axis represents more inappropriateness ratings, more severe punishment, and more morally egocentric moral judgments. The horizontal bar in the midst of the data represents mean value for the group.

(2) Details from planned comparisons carried out to decompose the main effect of group for each type of moral judgment.

- *Appropriateness judgments*

Bonferroni-corrected post-hoc comparisons to assess main effect of group revealed no significant differences.

- *Punishment judgments*

Bonferroni-corrected post-hoc comparisons to assess main effect of group revealed significant differences for all but attempted harm scenarios ($p = 0.280$, $BF_{10} = 0.996$). MS patients endorsed more severe punishment as compared to the control group not only for accidental (mean difference = 0.575, 95% CI [0.172, 0.977], $p = 0.024$, $BF_{10} = 7.131$) and intentional (mean difference = 0.671, 95% CI [0.163, 1.180], $p = 0.040$, $BF_{10} = 4.383$) harm cases, but also for neutral (mean difference = 0.561, 95% CI [0.186, 0.936], $p = 0.016$, $BF_{10} = 9.828$) cases.

To explore the source of this effect, we repeated this mixed ANOVA analysis separately for each independent variable that differed between two groups, viz. EC, PD, DDF, DIF, and EOT (see Table 1). All group differences remained significant in ANCOVAs with EC, PD, DDF, and DIF as covariates. But in ANCOVA with EOT as the covariate, all significant group differences vanished, as shown by Bonferroni-corrected post-hoc t -tests for neutral ($p = 0.344$, $BF_{10} = 1.378$), accidental ($p = 0.684$, $BF_{10} = 0.928$), attempted ($p = 0.244$, $BF_{10} = 1.278$), and intentional ($p = 0.196$, $BF_{10} = 1.747$) harm cases. In other words, once the differences in EOT were accounted for, MS patients as a group were no more punitive than the controls.

• Egocentricity judgments

Planned comparisons to explore the main effect of group for egocentricity judgments were carried out with Bonferroni-corrected comparisons. This analysis showed that groups comparisons were not significant for neutral cases and only marginally significant for accidental harm cases (neutral cases: $p = 0.744$, $BF_{10} = 0.249$; accidental harm: $p = 0.060$, $BF_{10} = 3.210$), but both group comparisons were significant for the negative belief (attempted harm: mean difference = 14.772, 95% CI [8.148, 21.396], $p < 0.001$, $BF_{10} = 634$; intentional harm: mean difference = 11.930, 95% CI [5.811, 18.049], $p < 0.001$, $BF_{10} = 111$).

As with punishment judgments, we carried out mixed ANCOVA separately for each covariate of interest (EC, PD, DDF, DIF, and EOT) to explore possible source of this effect. This analysis showed that significant group differences in egocentric moral judgments for attempted ($p = 0.175$, $BF_{10} = 1.083$) and intentional ($p = 0.314$, $BF_{10} = 0.714$) harm cases disappeared only when EOT was added as a covariate to the model.

Supplementary Text S6: Correlation analysis between independent and dependent variables

• Correlation analysis: Alexithymia and moral judgments

Correlation analysis between TAS scores and moral judgments about appropriateness and punishment ($\alpha = 0.05/4 = 0.0125$; for each judgment separately).

Variable	statistic	HC ($n = 38$)	MS ($n = 38$)
Punishment for neutral cases	ρ	0.056	-0.041
	p	0.737	0.809
Punishment for accidental harm	ρ	0.167	-0.102
	p	0.315	0.541
Punishment for attempted harm	ρ	-0.052	-0.137
	p	0.758	0.411
Punishment for intentional harm	ρ	-0.174	-0.134
	p	0.297	0.423
Appropriateness of neutral cases	ρ	0.191	0.063
	p	0.250	0.709
Appropriateness of accidental harm	ρ	0.123	-0.260
	p	0.463	0.115
Appropriateness of attempted harm	ρ	-0.169	0.034
	p	0.312	0.840
Appropriateness of intentional harm	ρ	-0.261	-0.291
	p	0.114	0.076

• Correlation analysis: Empathy and moral judgments

Spearman’s rho between IRI subscales for PT, EC, and PD and moral judgments about appropriateness and punishment for each group ($\alpha = 0.05/4 = 0.0125$; for each judgment separately). This analysis revealed that higher empathic concern was associated with higher punishment and inappropriateness ratings for attempted harm scenarios in controls, but not in MS population.

Variable	statistic	PT		EC		PD	
		HC (<i>n</i> = 38)	MS (<i>n</i> = 38)	HC (<i>n</i> = 38)	MS (<i>n</i> = 38)	HC (<i>n</i> = 38)	MS (<i>n</i> = 38)
Punishment for neutral cases	ρ	0.026	-0.066	0.286	0.041	0.042	0.165
	<i>p</i>	0.879	0.693	0.081	0.806	0.801	0.322
Punishment for accidental harm	ρ	-0.064	-0.044	-0.254	-0.034	-0.144	-0.070
	<i>p</i>	0.701	0.795	0.124	0.838	0.387	0.675
Punishment for attempted harm	ρ	-0.012	0.111	0.499	-0.040	0.295	-0.171
	<i>p</i>	0.943	0.509	0.001	0.812	0.073	0.304
Punishment for intentional harm	ρ	0.007	0.008	0.322	-0.143	0.129	0.108
	<i>p</i>	0.965	0.960	0.049	0.392	0.440	0.519
Appropriateness of neutral cases	ρ	-0.106	-0.158	0.121	0.126	0.067	0.134
	<i>p</i>	0.526	0.343	0.470	0.451	0.691	0.421
Appropriateness of accidental harm	ρ	-0.106	-0.207	-0.015	-0.188	-0.049	-0.192
	<i>p</i>	0.525	0.213	0.930	0.259	0.769	0.249
Appropriateness of attempted harm	ρ	0.092	-0.046	0.477	-0.046	0.400	0.119
	<i>p</i>	0.582	0.786	0.002	0.782	0.013	0.475
Appropriateness of intentional harm	ρ	0.045	-0.335	0.119	-0.391	-0.043	-0.184
	<i>p</i>	0.787	0.040	0.477	0.015	0.799	0.269

• *Correlation analysis: Clinical variables, alexithymia, empathy, and moral judgments*

Spearman's rho between clinical variables denoting disease severity with alexithymia scores, empathy subscales, and moral judgments. Since this analysis was restricted only to the MS group (*n* = 38), we did not correct for multiple comparisons to avoid too stringent thresholds for a small sample size.

The only significant correlated observed was between EDSS and personal distress, i.e. higher scores on self-reported disability was associated with greater distress in dealing with anxiety-provoking social situations.

variable	statistic	MSSS	EDSS	No. Of Relapses	Disease duration
DDF	ρ	.115	.187	.163	.160
	p	.491	.260	.330	.338
DIF	ρ	.171	.315	.262	.286
	p	.304	.054	.112	.082
EOT	ρ	-.032	-.071	.058	-.002
	p	.850	.674	.729	.990
TAS-20	ρ	.095	.176	.204	.213
	p	.569	.289	.220	.200
IRI_PT	ρ	.122	.149	-.136	.012
	p	.467	.372	.417	.942
IRI_EC	ρ	.168	.239	.113	.126
	p	.315	.148	.500	.452
IRI_PD	ρ	.247	0.326	-.022	.187
	p	.135	.046	.898	.260
Punishment for neutral cases	ρ	-.203	-.126	-.069	.278
	p	.221	.453	.680	.091
Punishment for accidental harm	ρ	-.289	-.278	-.234	.083
	p	.079	.091	.157	.619
Punishment for attempted harm	ρ	.154	.008	-.217	-.212
	p	.357	.963	.192	.200
Punishment for intentional harm	ρ	.180	.077	-.215	-.249
	p	.280	.647	.194	.131
Appropriateness of neutral cases	ρ	-.220	-.156	-.075	.148
	p	.185	.349	.653	.374
Appropriateness of accidental harm	ρ	-.302	-.289	-.090	.098
	p	.066	.079	.590	.557
Appropriateness of attempted harm	ρ	.035	-.020	-.267	-.045
	p	.836	.907	.106	.787
Appropriateness of intentional harm	ρ	.068	-.020	-.032	-.079
	p	.684	.907	.851	.639

Supplementary Text S7: Correlation analysis between wrongness and punishment judgments

Spearman’s *rho* between appropriateness and punishment moral judgments separately for healthy controls and MS patients for each condition. The difference in correlation values across groups was assessed using Fisher’s Z-test (2-tailed). The difference was considered significant if $Z > 1.96$.

Correlation pair for condition	statistic	HC ($n = 38$)	MS ($n = 38$)	Fisher's Z
neutral-neutral	ρ	0.738	0.66	0.641
	p	< 0.001	< 0.001	
accidental-accidental	ρ	0.658	0.802	1.317
	p	< 0.001	< 0.001	
attempted-attempted	ρ	0.654	0.577	0.520
	p	< 0.001	< 0.001	
intentional-intentional	ρ	0.692	0.413	1.726
	p	< 0.001	.010	

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