Adaptive Rewiring in Non-Uniform Coupled Oscillators

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Structural plasticity of the brain can be represented in a highly simplified form as adaptive rewiring, the relay of connections according to the spontaneous dynamic synchronization in network activity. Adaptive rewiring, over time, leads from initial random networks to brain-like complex networks, i.e., networks with modular small-world structures and a rich-club effect. Adaptive rewiring has only been studied, however, in networks of identical oscillators with uniform or random coupling strengths. To implement information processing functions (e.g., stimulus selection or memory storage), it is necessary to consider symmetry-breaking perturbations of oscillator amplitudes and coupling strengths. We studied whether non-uniformities in amplitude or connection strength could operate in tandem with adaptive rewiring. Throughout network evolution, either amplitude or connection strength of a subset of oscillators was kept different from the rest. In these extreme conditions, subsets might become isolated from the rest of the network or otherwise interfere with the development of network complexity. However, whereas these subsets form distinctive structural and functional communities, they generally maintain connectivity with the rest of the network and allow the development of network complexity. Pathological development was observed only in a small proportion of the models. These results suggest that adaptive rewiring can robustly operate alongside information processing in biological and artificial neural networks.

Keywords: evolving neural networks, neural oscillators, dynamical systems, complexity

Introduction

The anatomical connectivity of the brain network is shaped dynamically through structural plasticity (Butz et al., 2009). A variety of structural plasticity mechanisms serve the adaptive role of accommodating the functional connectivity, the mutual statistical dependencies between the network components (Avena-Koenigsberger et al., 2018; Rubinov et al., 2009). A common principle underlying these mechanisms has become known as *adaptive rewiring* (Gong & van Leeuwen, 2003, 2004; Papadopoulos et al., 2017). Adaptive rewiring implements the Hebbian principle of “what fires together, wires together” at the level of network dynamics (Bi & Poo, 2001; Hebb, 1949).

In networks of which the units may represent spiking model neurons (Kwok et al., 2007) or neural mass oscillators (Rubinov et al., 2009), adaptive rewiring involves estab-

lishing connections between dynamically synchronized units at the expense of currently non-synchronized ones. Over time, adaptively rewiring networks dynamically evolve into complex architectures, showing the characteristics of small-worldness, modularity, and the rich-club effect (Gong & van Leeuwen, 2003, 2004; Hellrigel et al., 2019; Rubinov et al., 2009). Since small-worldness (Sporns & Zwi, 2004), modularity (Meunier et al., 2010), and the rich-club effect (van den Heuvel & Sporns, 2011) are characteristics of large-scale brain connectivity, adaptive rewiring could represent a driving force for the dynamic evolution of brain connectivity structure.

Adaptive rewiring models typically are highly simplified models, of which the units are identical nonlinear oscillators with uniform connection weights and coupling strengths (Rubinov et al., 2009). These simplifications severely reduce the functionality of such models. To implement information processes in these networks, it is vital to enable non-identical oscillators and non-uniform connection strength. However, the question should be asked whether the properties of adaptive rewiring are robust to such non-uniformities. We therefore aimed to explore adaptive rewiring of coupled oscillators with non-uniform amplitude and coupling strengths and compare their evolution to that of uniform networks. The results will be crucial for the utility of adaptive rewiring in bi-

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ological and artificial neural networks with information processing functions such as pattern recognition and learning.

Adaptive Rewiring in Coupled Logistic Maps

Most adaptive rewiring studies have represented network activity by a model known as coupled logistic maps (Kaneko, 1992). This choice was reached based on a succession of abstractions. Mass activity was described by chaotic attractor dynamics, shown in Figure S1A of the Appendix (Breakspear et al., 2003). Further simplification was obtained via the Poincaré section, yielding the mapping in Figure S1B. This mapping, in turn, can approximately be described, minus the noise, by a logistic map (Figure S1C). Thus, the logistic map is the most straightforward possible abstract representation of neural mass activity (Rubinov et al., 2009). Note that in coupled map networks, the couplings provide noise perturbations to the oscillators, yielding mappings again more similar to Figure S1B.

The version of the logistic map used in our study is shown in Equation 1, in which x is a continuous variable in the range $[-1, 1]$ which is updated in discrete time t , and α is the *amplitude*. Logistic maps are known to exhibit universal dynamical properties (Feigenbaum, 1978). For certain regimes of α , the behavior of the logistic map converges to one or more limit-cycle attractors, but otherwise, it exhibits chaotic behavior. In these regimes, logistic maps produce deterministic bounded time series that, indeed, qualitatively resemble the oscillations of neural mass activity (see Figure S1).

$$x_{t+1} = 1 - \alpha x_t^2 \quad (1)$$

Because of the universal dynamics of logistic maps, networks of such simple maps may capture generic properties of interacting nonlinear systems (Kaneko, 1992). The logistic maps are coupled according to Equation 2, which describes how the activity of map i at time $t + 1$ as a function of its activity at the previous time step and the activity of maps coupled with it. In this equation, B_i denotes the set of units connected to unit i (its neighbors in the network), and the coupling strength ε_i sets the proportion to which the average activities of coupled units influence the activity of unit i (Hellrigel et al., 2019).

$$x_{i,t+1} = (1 - \varepsilon_i)(1 - \alpha_i x_{i,t}^2) + \frac{\varepsilon_i}{|B_i|} \sum_{j \in B_i} (1 - \alpha_j x_{j,t}^2) \quad (2)$$

Through the effect of the neighbors, the map activity of Figure S1C regains a noisy appearance more in line with Figure S1B. For efficiency of computation, we rewrite Equation 2 in matrix notation. For a network with N nodes, the activity of nodes at time $t + 1$ is calculated via Equation 3 (cf. Appendix for its derivation).

$$X_{t+1} = [\mathbf{1}_N - \alpha \odot X_t \odot X_t] \odot [\mathbf{1}_N - \varepsilon + (A_t \varepsilon) \oslash (A_t \mathbf{1}_N)] \quad (3)$$

In this equation, symbols $\odot$ and $\oslash$ denote Hadamard (i.e., element-wise) multiplication and division, respectively. The right-hand side of Equation 3 constitutes the vector form of the logistic map, in which α is the vector of amplitudes and X_t is the vector of node's activities at time t . The map is Hadamard-multiplied by a coupling term. In the coupling term, ε is the vector of coupling strengths, A_t is the connectivity matrix at time t , and $\mathbf{1}_N$ denotes a vertical unit vector of size N . In the coupling term, $A_t \varepsilon$ is Hadamard-divided by $A_t \mathbf{1}_N$, normalizing the former by the sum of the weights of the edges connected to each node. For binary networks, the term in the denominator counts the number of connections for each node. We consider only binary graphs for convenience. For adaptive rewiring in weighted networks, see Hellrigel et al. (2019).

Based on the network activity defined by Equation 3, adaptive rewiring takes the following form. After several updates to the network activity, a rewiring step is made. At each rewiring step, the connections of a random node are updated as follows: The node is disconnected from the neighbor most dissimilar in activity and is connected to the one most similar in activity amongst those it is presently unconnected to. Dissimilarity of two nodes at a given time is defined as the absolute value of the difference of their activity values. Note that although rewiring steps are local, the most dissimilar unconnected node is obtained through a global search. For algorithms using local or, instead, regional information for this purpose, see Jarman et al. (2017), Jarman et al. (2014).

All adaptive rewiring studies using coupled logistic maps have used fixed, uniform values for the amplitude parameters α and the coupling strengths ε . In logistic map networks with fixed architectures, these parameters have previously been allowed to vary in order to implement information processing functions. In a perceptual organization model (van Leeuwen et al., 1997), sensory input function was realized by local modulation of oscillator amplitudes. The presence of sensory input to certain units brought their amplitude parameter values down to impose a more stable regime on the oscillators. As a result, these oscillators showed a bias to become mutually synchronized. The synchronization was understood as representing perceptual grouping. Grouping preferences followed the Gestalt proximity principle and showed spontaneous switching in case of stimulation with bistable patterns.

In a memory model (van Leeuwen & Raffone, 2001), connectivity parameter values were locally incremented to represent the presence of a memory trace. This led to synchronization biases along the incremented connections, resulting in spontaneous rehearsal and relearning of stored pattern information. These examples demonstrate that coupled maps are capable of performing cognitive functions based on ongoing dynamics.

The question is whether the non-identical parameter settings involved in these functions would dovetail with adap-

tive rewiring. To answer this question, we partitioned networks into two subsets of units, a majority and a minority, each receiving a different parameter value for the amplitude α or the coupling strength ε . Fixing these parameters at different values involves a hardship test for adaptive rewiring. For instance, with one partition having increased coupling strength (or decreased amplitude), we might expect connections to be established preferentially within this partition and less within the other one. We may expect the opposite when coupling strength within the partition is decreased (or amplitude increased). In addition, we may expect fewer connections to be established between the partitions. This might induce network disintegration or otherwise interfere with the evolution of the network structure. On the other hand, if these perturbations fail to interfere with the emergence of complexity (i.e., small-worldness, modularity, and rich-club effect), adaptive rewiring could be used alongside pattern detection and learning in sparse neural networks. This has implications for the functioning of biological networks, as well as for the sparsification of artificial (i.e., deep) neural networks.

In what follows, the method section describes details of the composition and initialization of the models, the rewiring algorithm, and the qualitative and quantitative measures of network structures used to describe, characterize, and compare models. In the Results section, we describe our findings, mainly that non-uniformity of parameters is shown not to interfere with the evolution of brain-like structure while giving rise to distinguishable network structures suitable for cognitive functions. A discussion and several concluding remarks end the paper.

Method

Description of Networks

An unlabeled binary graph (or network) $G = (V, E)$ with n vertices (or nodes) and m edges (or couplings) is defined mathematically by a set of nodes $V = \{1, 2, \dots, n\}$ and a set of edges $E = \{(i, j) \in V^2 : i \text{ is coupled with } j\}$. E is also known as the adjacency list of the network. We use undirected graphs, i.e., $(i, j) \in E \Leftrightarrow (j, i) \in E$, and self couplings are not allowed, i.e., $(i, i) \notin E$. The size of the set V (i.e., the number of its members) is denoted by $|V|$. From E , we may construct the adjacency matrix A of G as a square matrix of the size $|V|$, the elements of which can take values of 0 or 1. The element on its i th row and j th column (i.e., a_{ij}) is equal to 1 if $(i, j) \in E$ and is zero otherwise. Since G is undirected, A is symmetrical around the main diagonal, and since there are no self connections, its diagonal elements are zeros.

We may partition V into two non-overlapping subsets, minority and majority, such that $|V_{\text{minority}}| = |V| - |V_{\text{majority}}|$ and $|V_{\text{minority}}| \ll |V_{\text{majority}}|$. The edges amongst members of these subsets form subgraphs within G are henceforth called *mi-*

nority and *majority partitions*. A third subgraph comprises of all of V but only edges between minority and majority nodes. Such a subgraph is called *interpartition*.

Dynamics on the Graph

To each $v_i \in V$, an activation value is assigned according to Equation 3. The corresponding parameters values, i.e., coupling strength ε and amplitude α , remain fixed in our model simulations. Models with identical parameter sets are called *families*. Five families of models are simulated, each with ten model instantiations, comprising a total of 50 model instantiations. Each network is run for 20 million iterations. All the simulations and analyses are conducted in R programming language version 3.6.0 (R Core Team, 2019) using computational resources provided by VSC (Flemish Supercomputer Center).

Parameter Setting and Initialization

In our models, all G s have $|V| = 300$ nodes and 5200 edges, a connectivity density providing robust evolution of small-world structure with uniform parameter setting (van den Berg et al., 2012). The model structure is initialized by randomly assigning $5200 \times 2 = 10400$ values “1” symmetrically to non-diagonal entries of A , and zeros to the remaining entries. Each node in the network is randomly and independently assigned an initial value, uniformly distributed between 0 and 1, i.e., $x_{i1} \sim \text{Unif}(0, 1)$.

Previous adaptive rewiring studies have been using values of α and ε in the ranges of $[1.7 - 1.9]$ and $[0.3 - 0.5]$, respectively (Gong & van Leeuwen, 2003; Hellrigel et al., 2019; van den Berg & van Leeuwen, 2004). Here the midpoints of these ranges, i.e., $\alpha = 1.8$ and $\varepsilon = 0.4$, are used for the parameters in the baseline (BL) condition. In the BL condition, all nodes have identical parameter values.

The same applies to the majority (250 nodes) of the other conditions. However, depending on the condition, the minority subset (i.e., the first 50) of nodes could have either lowered or increased values of either the α or ε parameters. As shown in Figure S2, higher values of the amplitude α tend to yield greater divergence in activity values; reducing the coupling parameter ε has a similar effect (Hellrigel et al., 2019). Conditions with lowered values of α are called less chaotic (LC), and those with increased values more chaotic (MC); conditions with lowered ε values are called sub-coupled (SC) and those with increased values hyper-coupled (HC). While keeping the parameters of the majority at the baseline level ($\alpha_{i \in 51:300} = 1.8, \varepsilon_{i \in 51:300} = 0.4$), five different combinations of parameters were assigned to the minorities, each combination called a “family”: The baseline family (BL; $\alpha_{i \in 1:50} = 1.8, \varepsilon_{i \in 1:50} = 0.4$), and the families with less chaotic minority (LC; $\alpha_{i \in 1:50} = 1.7, \varepsilon_{i \in 1:50} = 0.4$), more chaotic family (MC; $\alpha_{i \in 1:50} = 1.9, \varepsilon_{i \in 1:50} = 0.4$), sub-coupled minorities (SC; $\alpha_{i \in 1:50} = 1.8, \varepsilon_{i \in 1:50} = 0.3$), and

hyper-coupled minority (HC; $\alpha_{i \in 1:50} = 1.8, \epsilon_{i \in 1:50} = 0.5$). In the Results section, we identify model instantiations by the two capitals indicating their family, together with a serial number [1 – 10], e.g., BL7, MC10. The 10 model instantiations within each condition are run with different initializations, which are identical across conditions to allow matched comparison between families.

The Adaptive Rewiring Algorithm

A rewiring takes place after every 20 updates of the logistic maps, meaning that one million rewiring attempts are performed over the 20 million updates of the model. At each rewiring attempt, at time t , a node i is selected randomly from V , a vector of its distance from other nodes is calculated as $\mathbf{d}_{i,t} = |X_t - x_{i,t} \mathbf{1}_{|V|}|$, and another vector of similarities is defined as $\mathbf{s}_{i,t} = \mathbf{1}_{|V|} - \mathbf{d}_{i,t}$.

Using these vectors, we compute two vectors: $\boldsymbol{\delta} = [\delta_1, \dots, \delta_j, \dots, \delta_{|V|}]^T = A \mathbf{d}_{i,t}$ for the distances of node i from its neighbors; and $\boldsymbol{\sigma} = [\sigma_1, \dots, \sigma_j, \dots, \sigma_{|V|}]^T = (J_{|V|} - I_{|V|} - A) \mathbf{s}_{i,t}$ for the similarities of node i to its non-neighbors, where $J_{|V|}$ is a $|V| \times |V|$ unit matrix (with all elements equal to one) and $I_{|V|}$ is the identity matrix of size $|V|$ (with diagonal and off-diagonal elements equal to one and zero, respectively). The subtraction $J_{|V|} - I_{|V|} - A$ ensures the search for the most similar node takes place among non-neighbors. The most dissimilar neighbor and the most similar non-neighbor of node i , respectively denoted as ξ and ζ , are marked by finding the index of the maxima of $\boldsymbol{\delta}$ and $\boldsymbol{\sigma}$:

$$\begin{aligned} \xi &= \underset{j}{\operatorname{argmax}} \delta_j \\ \zeta &= \underset{j}{\operatorname{argmax}} \sigma_j. \end{aligned} \quad (4)$$

The rewiring is then changing the corresponding elements of the adjacency matrix A :

$$\begin{aligned} a_{i\xi} &= a_{\xi i} = 0 \\ a_{i\zeta} &= a_{\zeta i} = 1 \end{aligned} \quad (5)$$

Characterizing and Comparing Models

The state of each model at any given time t is described by adjacency matrix A_t (henceforth, “anatomical connectivity”), which is subject to adaptive rewiring, and the vector of activation values, X_t . A model’s “functional connectivity” at t (represented by the $|V| \times |V|$ matrix F_t) is defined by the momentary pairwise absolute differences of its node activation values.

Qualitative Description of Network Structures

Network structure can be qualitatively assessed by means of visual inspection of the wiring diagram or the adjacency matrix. Using the package *seriation* (Hahsler et al., 2008), the adjacency matrix is serialized by ordering rows and

columns according to the projection of the matrix on its first principal component. Seriation maximizes the visual identifiability of modules within the network.

Quantitative Measures of the Structure

After each rewiring attempt, we calculate six network connectivity measures (Costa et al., 2007): edge density, clustering coefficient, average path length, small-world index, modularity, and assortativity. Furthermore, we calculate an additional measure, namely, the rich-club coefficient, after the final rewiring. All measures are calculated separately for the entire graph and the three subgraphs (viz., the minority, majority, and interpartition subgraphs).

Edge Density. For a subset of nodes, this coefficient is the proportion of edges existing in a subgraph to the theoretical maximum number of edges possible in the same subgraph. For a subgraph with a subset of nodes $|V_s|$ and adjacency matrix A_s , this value is calculated as

$$ED = \frac{\sum_{ij} A_s}{|V_s| \cdot (|V_s| - 1)}. \quad (6)$$

Since the total number of edges remains the same during the adaptive rewiring, this coefficient gives an indication of how strongly each partition has attracted new nodes at every rewiring step.

Clustering Coefficient. This measure can be defined either locally or globally and gives an indication for the tendency of nodes to form clusters. We use the global clustering coefficient, which is defined as the number of closed triplets of nodes (the triplets of nodes that are all connected) divided by the number of connected triplets, either open (i.e., paths of length two) or closed (i.e., triangles). The numerator is equal to three times the number of triangles in the graph. Using linear algebra, the global clustering coefficient can be calculated formally from the adjacency matrix A via

$$C = \frac{3 \times \# \text{triangles}}{\# \text{triplets of connected nodes}} = \frac{\operatorname{Tr}(A^3)}{\sum_{ij} A_{ij}} A^2 - \operatorname{Tr}(A^2). \quad (7)$$

In this equation, $\operatorname{Tr}(A)$ is the trace of matrix A and is defined as sum of its diagonal elements, i.e., $\operatorname{Tr}(A) = \sum_i A_{ii}$.

Average Path Length. The average path length is the mean value of lengths of shortest paths between all pairs of nodes, as defined in Equation 4 for a network of size N , where d_{ij} is the length of the shortest distance between nodes i and j , and $d_{ij} = 0$ if there is no path between i and j . This measure, which we calculated using the *igraph* package (Csardi & Nepusz, 2006), gives an indication of how closely the nodes of a network are located from each other.

$$PL = \frac{1}{N(N-1)} \sum_{i \neq j \in V} d_{ij} \quad (8)$$

Small-world Index. This measure quantifies the degree to which a graph shows the optimal combination of local and global connectedness, known as *small-worldness* (Watts & Strogatz, 1998). It is defined as the proportion of clustering coefficient to average path length for a network,

$$SW = \frac{C}{PL}. \quad (9)$$

The small-world index is often normalized by $SW_0 = \frac{C_0}{PL_0}$, the expected small-world index of a random network of the same size and density as the network in question. C_0 and PL_0 are, respectively, the expected clustering coefficient and expected average path length in such a random network. Thus, the normalized small-world index is given by

$$SW_{norm} = \frac{SW}{SW_0} = \frac{\frac{C}{PL}}{\frac{C_0}{PL_0}}. \quad (10)$$

The normalized index makes the comparison of networks with different sizes and densities possible. Since all networks modeled in this study start off with random networks of equal size and density, for computational reasons, the non-normalized small-world index (Equation 9) is calculated and reported.

Modularity. Modularity of a network, as proposed by Newman (2006) and denoted by Q , is a measure of to what extent the nodes tend to form interconnected communities isolated from the other nodes of the graph. More precisely, for a network of size N (with the theoretical maximum number of edges $m = \frac{N(N-1)}{2}$) and adjacency matrix A , modularity is defined as

$$Q = \frac{1}{2m} \sum_{i,j} [M_{ij} - \frac{k_i k_j}{2m}] \delta(c_i, c_j). \quad (11)$$

In this equation, k_i and k_j are, respectively, degrees of nodes i and j . $\delta(c_i, c_j)$ is the Kronecker delta function, which is equal to one if nodes i and j have the same label and zero otherwise. The term in square brackets is the difference between the actual number of edges between nodes i and j and the expected number of edges between them. Hence, according to Equation 11, modularity equals to the sum of these differences for the nodes within communities, normalized by the theoretical maximum number of edges in the network.

This measure requires an a priori labeling of nodes that defines the communities to which the nodes belong. A variety of algorithms have been suggested to discover module, or communities, within a network so that the value of Q is maximized (for a review, cf. Zhang et al., 2018). The communities discovered by these algorithms can thus be used as labels for calculating modularity of the network. In line with Clauset et al. (2004), we use the fast greedy algorithm to optimally detect communities and thus calculate the modularity based on community membership of the nodes. The *igraph*

package was used to automatically detect communities and calculate modularity.

Assortativity. The assortativity coefficient is a measure of homophily in networks that indicates the preferences of nodes to connect to alike nodes. The likeness can be imposed externally, e.g., by assigning categories to the nodes using labels (known as nominal assortativity), or by internal criteria such as node degrees (degree assortativity). Degree assortativity is defined as the Pearson correlation coefficient of degrees of connected nodes, thus taking values in the range $[-1, 1]$.

To give a formal definition, let p_k be the probability that a randomly chosen node has degree k . It can be shown that the degree distribution for a node connected to a randomly selected edge l is thus proportional to $k p_k$. Then, a quantity for “remaining degree” is defined as $q_k = \frac{(k+1)p_{k+1}}{\sum_j j p_j}$ which is the normalized distribution of remaining degree for the nodes connected to l . Finally, the joint probability of remaining degrees of the nodes at both ends of l is denoted by e_{jk} . For an undirected network $e_{jk} = e_{kj}$ and its marginal distribution is $\sum_j e_{jk} = q_k$. Having the variance of remaining degree as $\sigma_q^2 = \sum_k k^2 q_k - [\sum_k k q_k]^2$, the degree assortativity can be calculated as shown in Equation 12 (Newman, 2003).

$$r = \frac{1}{\sigma_q^2} \sum_{i,j} j k (e_{jk} - q_j q_k) \quad (12)$$

Rich-club Coefficient. This coefficient quantifies the tendency of nodes with higher than a certain degree to connect to each other. More formally, as Equation 13 shows, the coefficient is equivalent to the edge density of the subgraph of the network where the nodes with lower degrees than the cut-off value k are removed,

$$RC(k) = ED_{\geq k} = \frac{\sum_{i,j} A_{\geq k}}{N_{\geq k}(N_{\geq k} - 1)}. \quad (13)$$

Since this coefficient is a function of club size k , it is hard to visualize its evolution over time for all possible values of k . Hence, the values of this coefficient were only plotted for the final state of the networks.

The absolute value of the rich-club coefficient is hard to interpret and is not comparable among networks of different sizes, densities, and degree distributions. Hence, this coefficient is often normalized by the average rich-club coefficient of random networks of the same size with similar degree sequence. For each model (and its minority and majority partitions) we simulate 200 such networks, for each club size k , and normalized the values of the rich-club coefficients of the network by the average rich-club coefficient among the random networks. Moreover, for each k , we statistically tested whether the non-normalized RC is significantly larger than the rich-club coefficients of randomly generated networks.

Since the distribution of RC s of random networks were non-Gaussian, the conventional 1-sample t-test was not applicable. Instead, we performed 1-sample Wilcoxon signed rank test and set $\alpha = 0.01$ as the significance level for the p-values.

For a certain k , a normalized rich-club coefficient larger than one indicates that nodes with degree k tend to connect to the “rich” nodes (i.e., those with degrees equal to or greater than k), thus forming “rich clubs.” Conversely, $RC_{norm}(k) < 1$ implies that the nodes with degree k show a tendency to connect to nodes with lower degrees. Finally, $RC_{norm}(k) = 1$ suggests that nodes with degree k show no preference to connect to nodes with lower or higher degrees.

Investigating the Resemblance Between Models

To compare network families with each other, we assume, in line with Berlingerio et al. (2012), that the structural information embedded in networks can be summarized by the distributions of local network measures. Comparison of networks is thus reduced to comparing these distributions. To obtain measures of distributional distances, we use the NetSimile method (Berlingerio et al., 2012) and the Heller–Heller–Gorfine (HHG) algorithm (Heller et al., 2013), both of which are discussed in the Appendix. In short, NetSimile indicates the degree of dissimilarity between the distributions attributed to the networks being compared via comparing signature vectors that encapsulate the distributions by their summary statistics.

HHG, on the other hand, provides p-values for a test of independence among the distributions; a small p-value derived from HHG (e.g., below the threshold of $\alpha = 0.05$) provides evidence to reject the null hypothesis of distributional independence. Thus, loosely speaking, HHG p-value can be regarded as an indicator for dissimilarity; a higher value of this measure entails a smaller “resemblance” (or dependence) between the networks. Nevertheless, interpreting HHG p-values as such measure is somewhat unorthodox and is hardly meaningful unless used along with another dissimilarity measure such as NetSimile.

We first make pairwise comparisons among the 1225 unique pairs of model instantiations at their final state after 1 million rewiring attempts. Having quantitative measures for dissimilarities among the networks, we quantify the within-family resemblances and between-family contrasts among the models. Finally, we define a score for family distinction in order to compare how families vary with respect to this measure.

Family Resemblances and Differentiations. The outcomes of pairwise comparison of the networks using NetSimile and HHG were stored in four 50×50 matrices of form $Dissimilarity_M^G$ for networks G (either A or F , respectively for anatomical and functional connectivities) based on method M (either NetSimile or HHG). To ease the visual comparison of these measures, the matrices of NetSimile dis-

tances, i.e., $Dissimilarity_{NetSimile}^A$ and $Dissimilarity_{NetSimile}^F$, were normalized by the highest value in each matrix so their values range from zero to one. Then, within- and between-family contrast aggregate scores were calculated by averaging the elements of dissimilarity matrices that belong to the families being compared as shown in Equation 14:

$$Contrast_M^G(f_i, f_j) = \frac{1}{10 \times 10} \sum_{i \in f_i} \sum_{j \in f_j} (Dissimilarity_M^G)_{ij}. \quad (14)$$

Finally, a differentiation score was calculated for each family to quantify the degree to which models belonging to family f_i resemble each other and, at the same time, diverge from the members of other families, via

$$Differentiation^G(f_i) = \frac{1 - Contrast_{NetSimile}^G(f_i, f_i)}{\frac{1}{4} \sum_{i \neq j} [1 - Contrast_{NetSimile}^G(f_i, f_j)]}. \quad (15)$$

In this equation, the numerator is the within-family resemblance of networks belonging to family f_i . The denominator is the mean of the between-family resemblance of f_i to other families. This score will be larger than one if family f_i differentiates from other families. For the lack of straightforward interpretation, this score was not calculated for HHG outcomes.

Results

Network Structures

Using the *seriation* and *igraph* packages, we plot the raw (unserialized) and ordered (serialized) adjacency matrices and the wiring diagrams of the networks in their final state. In the plots, the minority and majority nodes are colored sky blue and pink, respectively. In both matrix visualizations and wiring diagrams, the within-minority and within-majority edges are colored blue and red, respectively. The inter-partition edges, connecting nodes of minority subset to nodes of majority, are colored green. Although there are variations among models, either within- or between-families, in all models (except for the MC2, MC3, SC1, and SC3; see below), several densely coupled sets of nodes, i.e., modules, emerged. These modules can be identified as squares in the serialized adjacency matrices. Figure 1 shows two representative networks per family. As evident in the plots, the networks manifest a range of different structures. Yet, similarities can be observed among models belonging to the same families.

The baseline models (BL2 and BL8 in Figure 1) typically include three densely coupled modules and a few larger, sparser sets of nodes. The modules are not isolated from the rest of the network, as inter-modular edges keep them connected to other nodes. The unserialized adjacency matrices

Figure 1

Network structures of representative models. Each panel shows the unserialized (top left) and serialized (top right) adjacency matrices, and the graph representation (bottom) of the structural connectivity at the last rewiring step. The within-minority, within-majority, and interpartition edges are colored blue, red, and green, respectively. In the graph representation, the minority and majority nodes are colored sky blue and pink, respectively.

show that the density of edges is relatively uniform over subsets of nodes.

In the HC family, wherein the minority nodes have relatively higher coupling strengths, the edge density is higher in the minority subgraph. Moreover, the HC family networks have more distinct modules than other families. HC2 (Figure 1), for instance, only has two modules, both of which are highly connected. Higher edge density in the minority subset and highly distinct modules are also observed for the LC family, which has lower amplitude in the minority nodes (LC5 and LC8, Figure 1). The similarity between the HC and LC models was expected, as the nodes with lower amplitudes and higher coupling strengths can synchronize more easily. However, it is worth noting that the effect is not limited to the minority nodes; highly connected modules also emerge among the majority nodes.

In the MC family, the edge density in the minority partition is reduced. Moreover, the higher level of amplitude for minority nodes resulted in highly connected modules among the majority nodes (see MC4 in Figure 1). In the SC family, the edge density of the minority (which had lower coupling strengths) is lower than that of the majority. Additionally, the lower coupling strength of the minority prevented minority nodes from forming modules, and they were absorbed into modules formed mainly by the majority nodes. See SC5 and SC7 in Figure 1.

Network Statistics

The evolution of clustering coefficient, modularity, edge density, small-world index, degree assortativity, and average path length were plotted for the structural network connectivity of all models and their sub-networks (viz., within-minority, within-majority, and interpartition). Figures 2–6 show these plots for the minority, majority, and whole networks, grouped per family. Furthermore, the normalized rich-club coefficient of the final states of the minority, majority, and whole networks are plotted in Figure 7.

Let us first consider the evolution of network statistics for the whole network. As evident in the plots, modularity, clustering coefficients, and small-world index of all models increase noticeably after 60 thousand rewiring attempts (roughly by a factor of 4.7, 5.3, and 4.6, respectively) and degree assortativity approaches 0.53. Meanwhile, the average path length increases no more than 15%. The network statistics fluctuate afterward, and the fluctuation is more pronounced for degree assortativity. The majority subnetworks show similar trends. In the minority subnetworks—although with higher fluctuations and instability in HC, LC, and SC—we observe an increase in these statistics after 60 thousand rewiring attempts. The network statistics of minority subnetworks of MC and BL families resemble those of the whole networks in the same families. Yet, clustering coefficient and small-world index show higher growth in the minority sub-

Figure 2

Evolution of network statistics in the baseline (BL) condition for the whole network and majority and minority subgraphs.

networks (more than 6- and 7-fold, respectively). The exception to these improvements is the modularity, which increases only moderately (roughly 35%) in this subnetwork. Given the finite number of nodes in this partition, this was not unexpected; a larger subnetwork would likely have shown more improvement in modularity as well. The edge densities in the minority subnetworks, although unstable, are often large, roughly 2.4 times that of the whole network. In line with elevated edge density, average path length within the minority subgraph drops by almost 7%. This means that the minority nodes tend to connect more strongly to each other than to other nodes. Tables 1–3 summarize the mean and standard deviation of changes in network statistics after 60 thousand rewiring attempts for the whole network and the majority and minority subnetworks.

Figure 7 shows the normalized rich-club coefficient $RC_{norm}(k)$ of the minority, majority, and whole networks,

Table 1

Means and standard deviations (between parentheses) of network statistics of the whole graph for different conditions.

Condition	Clustering coefficient	Average path length	Small-world index	Modularity	Assortativity	Edge density
HC	5.33(1.05)	1.15(0.05)	4.62(0.80)	4.69(0.85)	0.52(0.22)	1(0.02)
MC	5.34(1.02)	1.15(0.05)	4.63(0.78)	4.7(0.81)	0.53(0.22)	1(0.01)
BL	5.32(1.05)	1.14(0.05)	4.62(0.80)	4.68(0.84)	0.53(0.22)	1(0.02)
LC	5.35(1.03)	1.15(0.05)	4.63(0.79)	4.7(0.82)	0.53(0.22)	1(0.02)
SC	5.35(1.01)	1.15(0.05)	4.64(0.78)	4.69(0.82)	0.53(0.22)	1(0.01)

Note: Means are from 60 thousand to one million rewiring steps, normalized by those of 100 random networks, except degree assortativity. BL: baseline; LC: less chaotic; MC: more chaotic; SC: sub-coupled; HC: hyper-coupled.

Table 2

Means and standard deviations (between parentheses) of network statistics of the majority subgraph for different conditions.

Condition	Clustering coefficient	Average path length	Small-world index	Modularity	Assortativity	Edge density
HC	5.49(1.13)	1.17(0.06)	4.64(0.85)	4.45(0.84)	0.52(0.18)	1.01(0.12)
MC	5.51(1.08)	1.17(0.06)	4.66(0.82)	4.47(0.8)	0.52(0.18)	1(0.11)
BL	5.49(1.12)	1.17(0.06)	4.64(0.85)	4.44(0.83)	0.52(0.18)	1.01(0.12)
LC	5.52(1.11)	1.18(0.06)	4.66(0.84)	4.47(0.81)	0.53(0.18)	1(0.12)
SC	5.53(1.08)	1.18(0.06)	4.67(0.83)	4.46(0.8)	0.54(0.18)	1(0.12)

Note: Means are from 60 thousand to one million rewiring steps, normalized by those of 100 random networks, except degree assortativity. BL: baseline; LC: less chaotic; MC: more chaotic; SC: sub-coupled; HC: hyper-coupled.

Table 3

Means and standard deviations (between parentheses) of network statistics of the minority subgraph for different conditions.

Condition	Clustering coefficient	Average path length	Small-world index	Modularity	Assortativity	Edge density
HC	6.24(1.71)	0.94(0.24)	7.16(3.19)	1.35(0.56)	0.53(0.32)	2.36(1.44)
MC	6.26(1.69)	0.93(0.24)	7.27(3.24)	1.35(0.57)	0.53(0.31)	2.42(1.49)
BL	6.18(1.72)	0.94(0.24)	7.07(3.03)	1.38(0.54)	0.52(0.32)	2.35(1.39)
LC	6.19(1.7)	0.93(0.23)	7.11(3.11)	1.39(0.52)	0.52(0.32)	2.34(1.38)
SC	6.2(1.67)	0.93(0.24)	7.17(3.25)	1.37(0.55)	0.51(0.31)	2.36(1.47)

Note: Means are from 60 thousand to one million rewiring steps, normalized by those of 100 random networks, except degree assortativity. BL: baseline; LC: less chaotic; MC: more chaotic; SC: sub-coupled; HC: hyper-coupled.

Figure 3

Evolution of network statistics in the less chaotic (LC) condition for the whole network and majority and minority subgraphs.

grouped by families, as a function of club size k . $RC_{norm}(k)$ above one (dashed line) indicates rich clubs. Values significantly larger than one ($p < 0.01$, based on 1-sample Wilcoxon signed rank test) are marked by solid circles. Despite differences among families, they all show significant rich-club structures for larger club sizes, both in the whole network and in the majority partition. Let us first consider the plots for the whole networks, depicted in the left column of this figure. The HC family has clubs of sizes $30 < k < 70$ (involving 10-23% of the nodes), most of them significant, with consistent RC_{norm} values among members. In the LC family, all models form rich clubs with $25 < k < 55$ (roughly 8-18% of the nodes), and some models form larger rich clubs up to the sizes of 65-85 (roughly 22-28% of the nodes). Other families, i.e., MC, BL, and SC, form clubs with sizes ranging from 30 to roughly 65 (10-22% of the nodes), with relatively higher (but less consistent) RC_{norm} values compared to HC.

Figure 4

Evolution of network statistics in the more chaotic (MC) condition for the whole network and majority and minority subgraphs.

In the majority partition (middle column of Figure 7), we consistently observe rich clubs larger than 25, involving 10% of the nodes within that partition. The HC family have rich clubs as large as 40 (16% of nodes), and other families form even larger rich clubs, with sizes spanning 45 to 60 (making up 18-24% of the nodes). In general, we observe more diverse values among models for the rich-club coefficients in the majority partition. Finally, in the minority partition (the right column of the same figure), we observe a relatively remarkable—and consistent—rich-club effect in the HC family ($12 < k < 30$; 24-60% of nodes), while other families have relatively larger rich-club coefficients for smaller club sizes ($5 < k < 20$; 10-40% of nodes). Some models of the LC family form much larger rich clubs with sizes up to 40 (i.e., 80% of the nodes).

Figure 5

Evolution of network statistics in the sub-coupled (SC) condition for the whole network and majority and minority subgraphs.

Incidental Losses

As can be observed in the evolution plots, four models (viz., MC2, MC3, SC1, and SC3) stop evolving before 10 thousand rewiring attempts. Visual inspection reveals that, at some point in their evolution, one node reaches maximum degree and is connected to all other nodes (cf. Figure 8). This indicates that these models reached the endpoint of a pathological development, something we chose not to prevent in our algorithm. Consequently, the rewiring algorithm naturally terminates, as division by zero looms in Equation 3. Terminated models failed to form any modules and were omitted from family-wise comparisons. We consider these incidental losses as a sign that robustness against perturbation is not always guaranteed for our models.

Figure 6

Evolution of network statistics in the hyper-coupled (HC) condition for the whole network and majority and minority subgraphs.

Family-wise Comparisons

NetSimile and HHG similarity measures were composed in new matrices wherein the lower triangle belongs to $Dissimilarity_{HHG}^N$ and the upper triangle belongs to $Dissimilarity_{NetSimile}^N$, i.e.,

$$Dissimilarities^N = \text{lower.tri}(Dissimilarity_{HHG}^N) + \text{upper.tri}(Dissimilarity_{NetSimile}^N). \quad (16)$$

The matrices of $Dissimilarities^N$ are plotted in Figure 9 as heat maps using the *ComplexHeatmap* R package (Gu et al., 2016).

The matrices of $Contrast^N$ are plotted as upper triangular matrices in Figure 10 using the *corrplot* R package (Wei & Simko, 2021). The cell colors, coded similarly to the heat maps, denote average contrast measures derived from

Figure 7

Normalized rich-club coefficients of the whole network after the last rewiring step, grouped by condition. Solid circles mark significant values.

the NetSimile algorithm while average HHG p-values (i.e., $Contrast_{HHG}^N$) are indicated in each cell. The HC family manifests the least within-family contrast. Based on the HHG test of multivariate independence, except for HC-HC and HC-LC family pairs, no conclusive evidence exists for distributional dependence among families. Finally, as can be seen in this figure, the within- and between-family NetSimile contrasts of both structural and functional connectivity networks show similar patterns. More specifically, both the structural and functional connectivity of HC-BL, HC-LC, BL-BL, BL-LC, and LC-LC, all share close contrast values compared to other family pairs. This is also the case for the BL-SC and LC-SC pairs.

The differentiation scores for structural and functional connectivity between families are plotted in Figure 11. Differentiation values above one (dashed line) imply that the within-family resemblance of network structures of family f_i is higher than the average resemblance of its members to

Figure 8

Network structures of terminated models. Panels and color coding are similar to those of Figure 1.

the members of other families. We observe elevated differentiation in both structural and functional networks of HC, BL, and LC. This measure is remarkably higher for the structural networks of the HC family while barely exceeding the threshold for the functional network of the same family.

Finally, in order to have both family resemblance and family differentiation in a single frame, we summarized their values in the graphs shown in Figure 12. In these graphs, individual nodes represent families of models. Edge color and size code between-family contrast and node color captures within-family contrast. The size of each node is proportional to the value of the differentiation score of its corresponding family. The families with $Differentiation^N(f_i) > 1$ are marked with asterisks. It can be noticed that in both structural and functional networks, the families with differentiation scores larger than one (i.e., HC, BL, and LC) have lower within-family contrast values. Moreover, the pairwise contrasts among these families (i.e., HC-BL, HC-LC, and BL-LC) are lower than any other family pair. This suggests that models with increased amplitude and decreased coupling strengths lead to structures less resembling the baseline fam-

Figure 9

Heat maps of pairwise dissimilarities of anatomical (top) and functional (bottom) networks. The upper diagonal elements show normalized dissimilarity measures derived from NetSimile algorithm, and the lower diagonal elements show HHG p -values. Model names and family assignments are indicated. Lower dissimilarity (hence higher similarity) measures are coded by brighter colors.

ily.

Discussion

We investigated the effect of non-uniformity of oscillator amplitude and connection strength on the evolution of network structure under adaptive rewiring. The adaptive rewiring was driven by synchronization in coupled logistic maps. The amplitude and coupling parameters of the oscillators govern their synchronization behavior. It has amply been

Figure 10

Heat maps of within- and between-family contrasts for anatomical (top) and functional (bottom) connectivities. The values within cells show the average HHG p -values of corresponding family-wise comparisons. Lower contrast measures are coded by brighter colors.

shown that networks evolve to complex, brain-like structures when these parameters were fixed to uniform values. We replicated this behavior for our baseline parameter setting. Additionally, we studied how non-uniform parametrization gets reflected in the evolving network structure and whether adaptive rewiring is robust to these perturbations.

To this aim, a minority subset of network units had either the amplitude reduced (LC) or increased (MC), or the connectivity reduced (SC) or increased (HC). These perturbations may be compared with perceptual input (modeled by increased or decreased divergence in node activities) or memory processes (modeled by varying degrees of cou-

Figure 11

Between-families differentiation scores of the anatomical and functional networks. Values above one (dashed line) imply above-average within-family resemblance compared to other families.

pling). In all these conditions, adaptive rewiring establishes the evolution of random networks into complex structures. These structures are modular and small-world networks and show a five-fold increase in the clustering coefficient, at the cost of an increase of less than 15% in average path length compared to the original random networks. Moreover, the models form rich clubs as large as 8-28% of network size. This evolution is in accord with previous adaptive rewiring studies. Adaptive rewiring, therefore, is generally robust against the symmetry-breaking perturbations of system parameters and thus can accommodate—and integrate—function-specific modules. Note, however, that robustness is not always guaranteed. In a small proportion of our models (four cases in total), we observed pathological development in the perturbed models, resulting in network evolution breakdown.

On top of these findings about the whole network, the subsets of our models also developed brain-like structures. More specifically, across families, the majority subnetwork shows improved properties akin to the whole network. The minority subnetwork of perturbed models—with an exception for modularity—manifests these properties to a much greater degree, especially in the relative sizes of its rich clubs. The latter observation is indicative of more specialized structures in the minority subnetwork as a direct result of heterogeneity in model parameters.

Despite the overall consistency in structural improvements of networks, we observed considerable variability among models, both in their evolution and the final states of structural and functional connectivity. These variations are partly due to the family-specific parameterization of models

Figure 12

Graph representation of family resemblance and differentiation in anatomical (top) and functional (bottom) connectivities of the fully evolved models. Edge color and size code between-family contrast and node color captures within-family contrast. Node size is proportional to the differentiation score of the corresponding family. The families with differentiation scores above one are marked with asterisks.

and partly to the random initialization of the models. We quantified the within- and between-family contrasts and defined a measure of family differentiation to score, on average, how well models belonging to one family distinguish from others.

We found that the baseline network family and families with increased coupling strengths or decreased amplitude (respectively, BL, HC, and LC) differentiate themselves most from other families, both in structural and functional connectivity. The differentiation score is the highest for the structural connectivity of the HC family. Moreover, we observe a relatively high resemblance between HC and LC for both structural and functional connectivity. Based on the HHG test, there is relatively strong evidence ($p = 0.02$) for distributional dependence between structural networks of HC

and LC. This means that elevated coupling strength and reduced amplitude have similar effects on both the structural and functional connectivity of adaptively rewiring oscillator models.

Altogether, perturbation to the parameters of coupled oscillators yields structural and functional differences that are fundamental for implementing cognitive functions in evolving networks. Moreover, for both the structural and functional connectivity, perturbation leads to differentiation from the baseline. Different perturbations show specific differentiations. From a cognitive neuroscience perspective, this implies that functional and structural differentiation can be used to identify functional components in a network, in accord with the use of structural and functional connectivity measures in neuroimaging.

Concluding Remarks

Further research may extend our findings in several ways. Firstly, the incidental occurrence of network evolution breakdown was unexpected. Technically speaking, the matrix algebraic implementation of coupled logistic maps is sensitive to minor computational errors such as undefined division for one node (which may stem from pathological development of models), as was the case for MC2, MC3, SC1, and SC3. A different implementation could have prevented breakdown by isolating the problematic node and proceeding the adaptive rewiring algorithm with the remaining network. However, rather than searching for an even more robust version of our algorithm, we choose to keep the current implementation for the sake of consistency with previous work, and as a warning signal that pathologies may arise as a by-product of neural network evolution. Future work may focus more specifically on such pathological development, in relation to developmental disorders in the human brain. This may require a systematic study of the probability distribution of network breakdown under a range of parametrization conditions, network sizes, and connectivity densities to find factors contributing to the pathological development of adaptively rewiring networks.

Secondly, a broader range of parameter variation conditions than currently imposed should be attempted. For instance, amplitude and coupling strength parameters could both deviate from baseline values simultaneously, either for the same subset of nodes or for two (overlapping or non-overlapping) subsets. Random and patterned deviations of parameters can be studied in large-scale systems to implement perceptual and memory functions. Ultimately, the aim is to have these functions implemented in a network that simultaneously maintains its optimal structure.

This study was limited to binary, unweighted, and undirected networks. The effect of non-uniform parameters of logistic maps can be studied, via systematic search, on weighted networks with various edge weight distributions,

akin to Hellrigel et al. (2019). For directed networks, see Rentzeperis et al. (2021).

We followed the strategy to provide the simplest possible model of brain structure and function in order to avoid stacking arbitrary assumptions. Simplification is inevitable in modeling; “all models are wrong—but some are useful.” (Attributed to George Box.) For those who consider the model to pose severe limitations on the generalization of our findings to neurobiological systems, more realistic neural mass models instead of coupled maps could be a viable solution, in particular, ones that have facilities for studying the effects of traveling waves and spatial embedding of the network topology (Calvo Tapia et al., 2020) on network evolution, or apply phase–amplitude–frequency coupling (Chehelcheraghi et al., 2016; Chehelcheraghi et al., 2017) to networks of neural mass activity (Deschle et al., 2021).

Finally, despite the current limitations, our study has set an essential first step in the development of adaptively rewiring networks capable of pattern recognition and learning. Showing the principled possibility of such networks offers scope for the study of the developing, functional brain as well as for applying adaptive rewiring to sparsify artificial neural networks (Gale et al., 2019).

Data and Code Availability

The reproducible scripts used in this study are available online on the study’s repository on the Open Science Framework (<https://osf.io/625d8/>). This repository also includes all model files of this study (500 files amounting to 101 GB of data, generated using VSC computational resources) and additional plots of various qualitative and quantitative network measures.

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Appendix

Derivations

Matrix-vector Realization of Coupled Logistic Maps

In the main body of the manuscript, we stated that the dynamics of a network of N coupled logistics map, each described by

$$x_{i,t+1} = (1 - \varepsilon_i)(1 - \alpha_i x_{it}^2) + \frac{\varepsilon_i}{|B_i|} \sum_{j \in B_i} (1 - \alpha_j x_{jt}^2) \quad (\text{S1})$$

may be re-written in matrix notation as

$$X_{t+1} = [\mathbf{1}_N - \boldsymbol{\alpha} \odot X_t \odot X_t] \odot [\mathbf{1}_N - \boldsymbol{\varepsilon} + (A_t \boldsymbol{\varepsilon}) \oslash (A_t \mathbf{1}_N)] \quad (\text{S2})$$

The equivalence of Equations S1 and S2 can be shown as follows. First, we may write the activities of all nodes at time $t + 1$ as a vector of size N , i.e.,

$$X_{t+1} = \begin{bmatrix} x_{1,t+1} \\ x_{2,t+1} \\ \vdots \\ x_{N,t+1} \end{bmatrix} = \underbrace{\begin{bmatrix} (1 - \varepsilon_1)(1 - \alpha_1 x_{1t}^2) \\ (1 - \varepsilon_2)(1 - \alpha_2 x_{2t}^2) \\ \vdots \\ (1 - \varepsilon_N)(1 - \alpha_N x_{Nt}^2) \end{bmatrix}}_{\text{II}} + \underbrace{\begin{bmatrix} \frac{\varepsilon_1}{|B_1|} \sum_{j \in B_1} (1 - \alpha_j x_{jt}^2) \\ \frac{\varepsilon_2}{|B_2|} \sum_{j \in B_2} (1 - \alpha_j x_{jt}^2) \\ \vdots \\ \frac{\varepsilon_N}{|B_N|} \sum_{j \in B_N} (1 - \alpha_j x_{jt}^2) \end{bmatrix}}_{\text{III}}. \quad (\text{S3})$$

We may write II as element-wise multiplication of two vectors, for which we use the Hadamard operator $\odot$:

$$\begin{aligned} \text{II} &= \begin{bmatrix} 1 - \varepsilon_1 \\ 1 - \varepsilon_2 \\ \vdots \\ 1 - \varepsilon_N \end{bmatrix} \odot \begin{bmatrix} 1 - \alpha_1 x_{1t}^2 \\ 1 - \alpha_2 x_{2t}^2 \\ \vdots \\ 1 - \alpha_N x_{Nt}^2 \end{bmatrix} \\ &= \left(\begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} - \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_N \end{bmatrix} \right) \odot \left(\begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} - \begin{bmatrix} \alpha_1 x_{1t} x_{1t} \\ \alpha_2 x_{2t} x_{2t} \\ \vdots \\ \alpha_N x_{Nt} x_{Nt} \end{bmatrix} \right) \\ &= (\mathbf{1}_N - \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_N \end{bmatrix}) \odot (\mathbf{1}_N - \begin{bmatrix} \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_N \end{bmatrix} \odot \begin{bmatrix} x_{1t} \\ x_{2t} \\ \vdots \\ x_{Nt} \end{bmatrix} \odot \begin{bmatrix} x_{1t} \\ x_{2t} \\ \vdots \\ x_{Nt} \end{bmatrix}) \\ &= (\mathbf{1}_N - \boldsymbol{\varepsilon}) \odot (\mathbf{1}_N - \boldsymbol{\alpha} \odot X_t \odot X_t). \end{aligned} \quad (\text{S4})$$

III can similarly be simplified using element-wise multiplication and division using Hadamard $\odot$ and $\oslash$ operators:

$$\begin{aligned} \text{III} &= \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_N \end{bmatrix} \odot \begin{bmatrix} \sum_{j \in B_1} (1 - \alpha_j x_{jt}^2) \\ \sum_{j \in B_2} (1 - \alpha_j x_{jt}^2) \\ \vdots \\ \sum_{j \in B_N} (1 - \alpha_j x_{jt}^2) \end{bmatrix} \odot \begin{bmatrix} \frac{1}{|B_1|} \\ \frac{1}{|B_2|} \\ \vdots \\ \frac{1}{|B_N|} \end{bmatrix} \\ &= \varepsilon \odot \underbrace{\begin{bmatrix} \sum_{j \in B_1} (1 - \alpha_j x_{jt}^2) \\ \sum_{j \in B_2} (1 - \alpha_j x_{jt}^2) \\ \vdots \\ \sum_{j \in B_N} (1 - \alpha_j x_{jt}^2) \end{bmatrix}}_{\text{III}} \oslash \underbrace{\begin{bmatrix} |B_1| \\ |B_2| \\ \vdots \\ |B_N| \end{bmatrix}}_{\text{IV}}. \end{aligned} \quad (\text{S5})$$

III is sum of activities of nodes connected to node i , and we may calculate III by left-multiplying the adjacency matrix A_t by the vector of activities of all nodes:

$$\text{III} = A_t(\mathbf{1}_N - \alpha \odot X_t \odot X_t). \quad (\text{S6})$$

This multiplication implies that III_i (the i th element of the vector) is equal to $\sum_j a_{ij}(1 - \alpha_j x_{jt}^2)$, in which a_{ij} is the element on row i and column j of A_t , which is 1 if nodes i and j are connected and 0 otherwise. In other words, this multiplication serves the selection of neighboring nodes and summation of their values.

In a similar manner, IV is a vector in which the i th element is the number of nodes connected to node i . Thus, this vector is equal to row-wise summation of A_t , which can be calculated using matrix multiplication via

$$\text{IV} = A_t \mathbf{1}_N. \quad (\text{S7})$$

Thus, given commutative, associative, and distributive properties of Hadamard operators, we have

$$\begin{aligned} \text{III} &= \varepsilon \odot [A_t(\mathbf{1}_N - \alpha \odot X_t \odot X_t)] \odot (A_t \mathbf{1}_N) \\ &= (\varepsilon A_t) \odot (\mathbf{1}_N - \alpha \odot X_t \odot X_t) \oslash (A_t \mathbf{1}_N) \\ &= (\mathbf{1}_N - \alpha \odot X_t \odot X_t) \odot (\varepsilon A_t) \oslash (A_t \mathbf{1}_N). \end{aligned} \quad (\text{S8})$$

Putting Equations S4 and S8 together and rearranging terms proves the equivalence:

$$\begin{aligned} X_{t+1} &= \text{II} + \text{III} \\ &= (\mathbf{1}_N - \varepsilon) \odot (\mathbf{1}_N - \alpha \odot X_t \odot X_t) \\ &\quad + (\mathbf{1}_N - \alpha \odot X_t \odot X_t) \odot (\varepsilon A_t) \oslash (A_t \mathbf{1}_N) \\ &= (\mathbf{1}_N - \alpha \odot X_t \odot X_t) \odot [(\mathbf{1}_N - \varepsilon) + (\varepsilon A_t) \oslash (A_t \mathbf{1}_N)]. \end{aligned} \quad (\text{S9})$$

Methods

Networks as Distributions

Network statistics can describe network structures either locally or globally. Local measures are suitable for node-

wise (or clique-wise) comparisons, while the global measures are aggregates of some local properties that provide summary statistics for the structure. Whereas local measures hardly lead to a holistic description of networks (as the nodes are usually described in isolation from each node) in their aggregation for deriving global measures, structural information is sacrificed. Therefore, neither local nor global measures are optimally suitable for the comparison of networks. A possible solution is to use the distribution of several local measures for network comparison.

Berlingerio et al. (2012) suggest characterizing each node i of the network with a seven-dimensional feature vector consisting of the following local measures that capture characteristics of the node and its first- and second-order neighbors:

1. $d_i = |N(i)|$, the number of neighbors, or degree ($N(i)$ is the set of first-order neighbors of node i , i.e., the nodes directly connected to i);
2. c_i , local clustering coefficient, which is the number of triangles connected to node i over the number of connected triples centered on node i ;
3. $\overline{d_{N(i)}} = \frac{1}{d_i} \sum_{j \in N(i)} d_j$, the average degree of node's neighbors;
4. $\overline{c_{N(i)}} = \frac{1}{d_i} \sum_{j \in N(i)} c_j$, the average clustering coefficient of the node's neighbors;
5. $|E_{ego(i)}|$, the number of edges in the *egonet* of node i ; the *egonet* of node i , referred to as $ego(i)$, is the subset of the network including node i , its first-order neighbors ($N(i)$), and the edges among $N(i)$;
6. $|E_{ego(i)}^\circ|$, the number of outgoing edges from the *egonet* of node i ; and
7. $|N(ego(i))|$, the number of neighbors of the *egonet* of node i .

Although more local features could be added to this vector, Berlingerio and colleagues have shown that these features suffice in practice for comparison of networks.

The NetSimile Method

The distribution of local features enables the comparison of their summary statistics. In NetSimile (Berlingerio et al., 2012), the feature distribution (which is a *nodes $\times$ features* matrix) is aggregated into a 35-dimensional "signature vector" consisting of five summary statistics for each feature: median, mean, standard deviation, skewness, and kurtosis. The comparison of networks is thus reduced to calculating distances (or similarities) of the signature vectors. NetSimile is superior to other methods of inferring network similarity, as its computational complexity grows linearly with network

size. More importantly, NetSimile allows the comparison of networks of different sizes.

Ranks and values of summary statistics characterize the overall shape of distributions and thus are highly diagnostic for their comparison (Berlengerio et al., 2012). Hence, the signature vectors are akin to ranked lists. It has been shown that the Canberra distance, defined in Equation S10 for two vectors V and W of size n , is an appropriate measure of dissimilarity for ranked lists (Jurman et al., 2009), as it is sensitive to small distances from zero and normalizes the pairwise distances of features by their absolute values. Moreover, Berlengerio et al. (2012) report that Canberra distance outperforms other distance measures in discriminating signature vectors, a desirable property of a dissimilarity measure for the task at hand.

$$\text{Ca}(V, W) = \sum_{i=1}^n \frac{|v_i - w_i|}{|v_i| + |w_i|}, \quad (S10)$$

We use this dissimilarity metric in the pairwise comparison of the signature vectors derived from the NetSimile algorithm. However, NetSimile does not allow hypothesis testing to infer significance levels for the distances. Berlengerio et al. (2012) suggest hypothesis testing for the independence of the distributions by pairwise comparison of the univariate distributions of the features and aggregating their p-values through averaging or choosing the maximum values. The authors report that neither Mann–Whitney nor Kolmogorov–Smirnov tests—which are nonparametric tests without any assumption for the distributions being compared—yield meaningful discrimination among networks. Notably, their approach of hypothesis testing ignores the multivariate dependencies amongst the features. Hence, we use another method to test the independence of distributions which is discussed below.

Hypothesis Testing for Similarities of Network Distributions

The significance tests used by Berlengerio et al. (2012) posit multivariate independence among features and lacks what the authors call “discrimination power.” To tackle this issue, one must use multivariate dependence tests. Since parametric dependence tests make assumptions about the distributions being compared, we used the Heller–Heller–Gorfine (HHG) nonparametric permutation test of multivariate dependence (Heller et al., 2013) implemented in the *HHG* R package (Brill et al., 2018). HHG is a consistent omnibus test for the null hypothesis of distributional independence, i.e., that the joint distribution of two multivariate random variables X and Y is equal to the multiplication of the marginal distributions of those variables. Equation S11

shows the null and alternative hypotheses:

$$\begin{cases} H_0 : F_{XY}(x, y) = F_X(x)F_Y(y) \\ H_1 : F_{XY}(x, y) \neq F_X(x)F_Y(y) \end{cases}, \quad \forall x, y. \quad (S11)$$

HHG has a reasonable computational complexity and uses norm distance matrices of the samples taken from X and Y separately. The technical details of this method are beyond the scope of this paper. In short, HHG iteratively forms hyperspheres in the joint space of $F_{XY}(x, y)$, and based on the implications of the null hypothesis, quantifies evidence against H_0 by likelihood ratio or Pearson’s Chi-square tests statistics over contingency tables. From these tests, one can drive permutation p-values that can be interpreted as evidence against the null hypothesis of independence of the distributions. Hence, the lower the p-value, the more evidence favoring the dependence of the distributions being compared. Loosely speaking, one can treat the p-values derived from HHG methods as a form of distributional dissimilarity; the lower the p-value, the more “similar” the distributions are to each other. This interpretation allows us to compare non-significant p-values as relative measures of resemblance.

The *hhg.test()* function in *HHG* package runs the test for a number of permutations on distance matrices of the samples in X and Y . It outputs four different permutation p-values based on sums or maximum values of likelihood ratio or Chi-square test scores of all 2×2 contingency tables. We let HHG run for 2000 permutations for each pairwise comparison and extracted permutation p-value for the maximum of likelihood ratio score statistics as it yielded higher discriminative power compared to other test statistics.

Figure S1

Dimension reduction of nonlinear neuronal dynamics. (A) Phase space attractor of a three-dimensional neural mass flow. This attractor is an illustration of the dynamics generated by the flow of a neural mass mode (cf. Breakspear et al., 2003). The dynamical variables represent the mean membrane potential of pyramidal (V) and inhibitory (Z) neurons, and the average number of open potassium ion channels (W). (B) Poincaré first return map from the same attractor (cf. Breakspear et al., 2003); this map captures key features of the neural mass flow, by following each trajectory from one intersection (V) of the attractor to the next ($P(V)$). (C) The quadratic logistic map that has the same unimodal topology as the neural mass Poincaré return map. While the logistic map lacks the “thickness” of the neural mass map, it is several orders of magnitude faster to compute, hence allowing more detailed quantitative analysis. Figure and caption adapted from “Symbiotic relationship between brain structure and dynamics,” by M. Rubinov, O. Sporns, C. van Leeuwen, and M. Breakspear, 2009, BMC Neuroscience, 10(1), 55. CC BY 2.0.

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Figure S2

Feigenbaum diagram of the values of the 200 draws of logistic maps (after a burn-in period of 4000 iterations). For three levels of amplitude used in this study (i.e., $\alpha = 1.7, 1.8$, or 1.9 ; vertical lines), logistic maps exhibit chaotic behavior, making them suitable dynamic elements in our models.