Investigating variability in morphological processing with Bayesian distributional models

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Abstract

We investigated the processing of morphologically complex words adopting an approach that goes beyond estimating average effects and allows testing predictions about variability in performance. We tested masked morphological priming effects with English derived ('printer') and inflected ('printed') forms priming their stems ('print') in non-native speakers, a population that is characterized by large variability. We modeled reaction times with a shifted-lognormal distribution using Bayesian distributional models, which allow assessing effects of experimental manipulations on both the mean of the response distribution ('mu') and its standard deviation ('sigma'). Our results show similar effects on mean response times for inflected and derived primes, but a difference between the two on the sigma of the distribution, with inflectional priming increasing response time variability to a significantly larger extent than derivational priming. This is in line with previous research on non-native processing, which shows more variable results across studies for the processing of inflected forms than for derived forms. More generally, our study shows that treating variability in performance as a direct object of investigation can crucially inform models of language processing, by disentangling effects which would otherwise be indistinguishable. We therefore emphasize the importance of looking beyond average performance and testing predictions on other parameters of the distribution rather than just its central tendency.

Keywords: RT distribution, distributional models, masked priming, visual word recognition, morphological processing

Introduction

Psycholinguistic research, and cognitive science more generally, is primarily concerned with estimating average effects for population samples in a given task. The underlying assumption is that human performance is largely homogeneous—across participants, items, or trials—and, consequently, average effects can be taken to reflect the cognitive mechanisms involved in performing the task. Hence, variability in performance has not traditionally been seen as a potential source of information about the structure of the cognitive system. Instead, it has been mostly ignored or dismissed as noise (Andrews, 2012). However, recent research has witnessed increasing interest in variability in language processing, with a growing consensus that this is a reflection of a flexible system, and therefore an intrinsic aspect of language processing mechanisms (see Amenta & Crepaldi, 2016; Kidd et al., 2018). This makes investigating variability crucially informative for building theoretical models of language processing.

The present work focuses on variability in morphological processing, that is, in processing complex words such as *player* [play][*-er*] or *played* [play][*-ed*], during visual word recognition. A widely employed technique in morphological processing research is the *masked priming* paradigm. In a masked priming experiment, participants are typically required to perform a lexical decision on a series of visually presented target words. These are preceded by so-called prime words, which are presented very briefly (around 50 ms) and are in turn preceded by a visual mask, thus preventing their conscious recognition (Baayen, 2014; Kinoshita & Lupker, 2004). Prime words can be morphologically related to the corresponding target, such as in *walked-WALK*, or completely unrelated in form or meaning, e.g., *kissed-WALK* (e.g., Feldman & Soltano, 1999; Rastle et al., 2000). When prime-target pairs are morphologically related, reaction times (RTs) are on average faster as compared to the unrelated condition, an effect known as ‘morphological priming’ (Amenta & Crepaldi, 2012). Morphological priming has been attributed to pre-activation of the target word during the processing of the morphologically related prime, due to decomposition into its morphological constituents—e.g., ‘walked’ → ‘walk’ + ‘ed’ (Rastle et al., 2004; Taft & Forster, 1975)—or, according to more recent accounts, to extraction of edge-aligned embedded words from letter strings—e.g., extracting ‘walk’ from ‘walked’ (Grainger & Beyersmann, 2017). Alternatively, morphological priming effects have been explained in terms of shared semantic and orthographic properties of prime and target, without positing an independent level of morphological representation (Baayen et al., 2011; Baayen & Smolka, 2020; Feldman, 2000).

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Masked priming effects have been reported for a series of different morphological phenomena from many typologically different languages (see Ciaccio et al., 2020, for a review). All of these studies have focused on the effect of prime presentation on *average* RTs, neglecting the fact that the same average RT may come from different RT distributions, possibly with different levels of variability (standard deviation) around the mean. In recent psycholinguistic studies from other domains than masked morphological priming, the standard deviation around average effects has been shown to systematically vary as a function of experimental manipulations (Chuang, Fon, et al., 2021; Tomaschek et al., 2020). This would raise the question as to whether different types of morphological phenomena, despite showing comparable priming effects on average RTs, are associated with different levels of variability around such average effects, and if so, why.

Besides providing a more thorough picture of the underlying data, investigating variability in morphological processing is also of theoretical interest. In lexical decision experiments, mostly involving non-native (L2) speakers, RTs have been shown to become less variable on a trial-by-trial basis when lexical representations become better established, for example with practice, and thus processing becomes more efficient (N. S. Segalowitz & Segalowitz, 1993; S. J. Segalowitz et al., 1998). Moreover, in novel word-learning studies, trial-level variability has been shown to increase during the process of establishing new representations (Solovyeva & DeKeyser, 2018). Variability in lexical processing has been therefore taken as a crucial index of processing automaticity, with a higher level of variability reflecting less efficient, or less automatic, lexical processing (N. S. Segalowitz, 2008). While several studies have looked at variability in performance in the context of second or artificial language learning, including some testing morphosyntax (e.g., Rodgers, 2011; for a review, see N. S. Segalowitz, 2008), none of them have specifically looked at online morphological processing.

By extending the account of variability in lexical processing proposed in N. S. Segalowitz and Segalowitz (1993) and subsequent studies to the domain of morphological processing, we may take the amount of variability in RTs in a masked morphological priming paradigm to reflect how efficiently, or automatically, speakers are able to activate the morphological form presented in the prime (e.g., ‘printer’ or ‘printed’) and, from that form, activate the stem (‘print’). Note that a ‘standard’ masked priming effect on average RTs can only go as far as providing some evidence for lexical activation of a stem from the morphologically complex prime, while it does not tell us anything about how efficiently this process operates on a trial-by-trial basis. This way, analyses of *variability* around *average* RTs would crucially complement standard RT analyses of masked priming data.

A population that lends itself particularly well to the investigation of trial-level variability in language processing are L2 speakers. Performance in an L2 is characterized by larger heterogeneity than in native language, both within and across individuals (Bialystok & Hakuta, 1999; Hopp, 2013; White, 2003). When it comes to morphological processing, results from L2 speakers tend to be more variable across studies than results from native speakers. However, there is at least some indication that variability may also be selective, or at least more enhanced, for some types of morphologically complex words as opposed to others. While masked morphological priming studies with L2 speakers have generally

reported robust priming effects from derived forms (Ciaccio & Clahsen, 2020; Diependaele et al., 2011; Heyer & Clahsen, 2015; M. Li et al., 2017), the size of priming effects with inflected words tends to be smaller or to substantially vary across studies (Jacob et al., 2018; Veríssimo et al., 2018; Feldman et al., 2010). This is in line with previous work suggesting persistent vulnerabilities in L2 speakers for inflectional morphology, especially with respect to the ability of consistently and reliably accessing inflected forms and morphosyntax (Blom et al., 2006; Hopp, 2013; White, 2003). Following the approach to variability by N. S. Segalowitz and Segalowitz (1993), the inability of consistently accessing inflected forms should not only impact the size of an average experimental effect of an L2 group, but it should also be reflected in more trial-by-trial performance variability. We therefore took the inflection-derivation dichotomy in an L2 as a particularly good test case for investigating RT trial-level variability in morphological processing, and whether this varies depending on the type of complex words being processed.

The present study

The experiment reported below contrasted the processing of English inflected and derived forms using the masked priming paradigm. We investigated the effect of these two types of morphologically complex words on the visual recognition of their bases. Specifically, we assessed priming effects elicited by inflected *-ed* past-tense forms (e.g., *printed*) and derived *-er* nominalizations (e.g., *printer*) on lexical decision times to the same target stems (e.g., *print*). The two conditions were created as to be comparable in a range of morphological and non-morphological properties: (a) both *-ed* and *-er* are regular, productive, and semantically transparent suffixes; (b) the two types of morphological primes were equivalent in their amount of orthographic overlap with their targets; and (c) priming conditions were well matched in other lexical measures (see below). Crucially, priming effects were estimated not only on average response speed but also on the variability of responses.

We made use of Bayesian *distributional models* to assess the effects of morphologically related primes. In distributional models, estimates refer not only to the mean of responses in different experimental conditions but can also describe additional features of the response distribution, for example, its variability. Moreover, these models allow the different parameters of a distribution to depend on a set of explanatory variables. For example, RTs can be estimated to be shorter or longer in terms of their central tendency, but also narrower (less variable) or wider (more variable), depending on a given predictor (Bürkner, 2018; Kneib & Umlauf, 2017). As a result, distributional models provide better estimates of differences between means (because they relax the common assumption of equal variances), but more importantly, they allow going beyond the mean in order to draw inferences about how the whole shape of a response distribution is affected by the experimental manipulations (see Balota et al., 2008; Balota & Yap, 2011).¹

Another advantage of Bayesian models is that they allow fitting virtually any kind of

¹ Effects on both the mean and the standard deviation of a dependent variable can also be tested with generalized additive mixed models [GAMMs; see Baayen et al. (2017)] with the *mgcv* package (Wood, 2017) in R. We therefore provide an alternative analysis of our data with GAMMs in the Supplementary Materials S3. We thank Harald Baayen for this useful suggestion.

distribution in a straightforward way (Nicenboim et al., 2016). Here, we modeled RTs as a ‘shifted-lognormal’ distribution, which is described by three parameters (see Figure 1 for examples): (a) *mu*, the mean of (normally distributed) RTs in the log scale, (b) *sigma*, the standard deviation of (normally distributed) RTs in the log scale, and (c) *shift*, which moves the whole RT distribution to the right. The mean *mu* is a ‘location’ parameter, expressing central tendency, and can be taken as an index of difficulty (Wagenmakers & Brown, 2007): experimental manipulations that slow down responses, for example, can be empirically described as increasing *mu*, which disperses the RT distribution in the direction of longer reaction times. The standard deviation *sigma* is a ‘scale’ parameter, which stretches or squeezes the RT distribution around the same center. In this way, the *sigma* parameter captures variability of responses, and importantly, does so independently of their central tendency, that is, over and above the effects of condition difficulty (Wagenmakers & Brown, 2007). Finally, the shift is an estimated quantity in milliseconds that is added to the RT distribution and constitutes a lower bound for responses; this parameter may correspond to more peripheral (visual or motor) aspects of processing (Logan, 1992; Rouder, 2005)². One advantage of including a shift parameter (relatively to assuming a lognormal distribution or to an analysis of log-RTs) is that shifted-lognormal models can more closely approximate the shape of RT distributions, and thus better satisfy the assumptions of linear models.

Example of shifted-lognormal distributions

Figure 1

Examples of shifted-lognormal distributions, with varying *mu* (location), *sigma* (scale), and shift parameters. The *mu* and *sigma* parameters are expressed in log-milliseconds, whereas the shift is expressed in milliseconds. In a shifted-lognormal distribution, *mu* and *sigma* are the mean and standard deviation of the normal distribution of $\log(RT - \text{shift})$. Conversely, an interpretable back-transformation of *mu* to the millisecond scale is given by $\exp(\mu) + \text{shift}$, which is equal to the median of RTs.

² We here refer to common interpretations of the different parameters of RT distributions. However, note that whether these parameters neatly map onto specific cognitive processes has also been a matter of debate (see Matzke & Wagenmakers, 2009).

We made use of such distributional models to assess priming effects on both the *mu* and *sigma* parameters of the shifted-lognormal distribution of lexical decision RTs. In line with some of the previous morphological priming studies with L2 groups (Jacob et al., 2018; Kirkici & Clahsen, 2013; Veríssimo et al., 2018), we expected a difference between derivational and inflectional priming on the *mu* (mean) of log-RTs. Specifically, derived forms should produce facilitation effects on the recognition of their constituent stems, but the masked priming effects elicited by inflected forms may be smaller in magnitude or absent. Additionally, if L2 speakers show particular difficulties with consistently accessing and decomposing inflected forms in written word recognition, then the presentation of inflected primes may produce more inconsistent benefits, possibly leading to more variable lexical decision responses. In that case, inflected primes are expected to increase the *sigma* (standard deviation) of log-RTs in comparison to derived (and possibly unrelated) primes.

Method

Participants

Sixty-nine intermediate to advanced non-native speakers of English (54 women; 15 men) took part in the experiment in exchange for payment or course credits. Their mean age was 26.09 years (SD = 5.27, range = 18–37). All participants started learning English after the age of 4 (mean age of acquisition = 8.71, SD = 1.95, range = 4–13; 3 NAs). Participants were all native speakers of German, three of whom additionally spoke Russian as a native language. They all lived in Germany at the time of testing. All participants reported reading in English to some extent in their daily lives (mean use of English, as compared to other languages = 32.14%, SD = 18.10, range = 2–80%). Skill in English was tested by means of a 50-item multiple-choice grammar test adapted from the Oxford Placement Test 1 (Allan, 2004). Participants' mean score was 38.62/50 (SD = 5.91; range = 21–48). This corresponds to a proficiency level roughly ranging between B2 and C2 of the Common European Framework of Reference for Languages (Council of Europe, 2001), i.e., from upper intermediate to highly proficient. All participants additionally took the English version of the LexTALE test (Lemhöfer & Broersma, 2012), a standardized vocabulary test consisting of an un-speeded visual lexical decision task. The LexTALE score is a percentage score, calculated as the percentage of correct responses corrected for the proportion of existing and non-existing words in the test. The group achieved a mean score of 78.91% (SD = 9.28; range = 61.25–98.75%, again roughly corresponding to B2 to C2 level). Prior to testing, all participants signed a written consent.

Materials

The experiment included 102 English monomorphemic verbs used as targets (e.g., *print*). These were preceded by their *-ed* past-tense form (e.g., *printed*) as the inflected prime, their *-er* nominalization (e.g., *printer*) as the derived prime, or by an unrelated prime. Unrelated primes were dissimilar in form and meaning from their corresponding targets; half of them were *-ed* inflected forms and half of them were *-er* derived words. The distributions of word-form frequency, lemma frequency, and length (in letters) of the three prime types (inflected, derived, unrelated) were kept as similar as possible. Word-form

and lemma frequency³ were extracted from the SUBTLEX-UK database (van Heuven et al., 2014) and are provided in the Zipf scale, which approximately spans from 1 to 7; values below 3 indicate relatively low frequency, while values above 4 indicate high frequency. Item characteristics are provided in Table 1.

Table 1

Summary of the item characteristics (mean, SD, range)

Item type	Length (letters)	Word-form frequency	Lemma frequency
Target	5.09 (1.24) 3–8	3.93 (0.62) 1.9–5.37	4.04 (0.53) 2.55–5.21
Derived prime	6.73 (1.07) 4–9	3.03 (0.64) 1.3–4.22	3.2 (0.65) 1.6–4.46
Inflected prime	6.73 (1.07) 4–9	3.29 (0.73) 1.17–4.76	3.88 (0.58) 2.23–5.11
Unrelated prime	6.73 (1.07) 4–9	3.24 (0.73) 1.3–4.73	3.66 (0.84) 1.3–5.57

The 102 experimental targets and their corresponding inflected, derived, and unrelated primes were distributed across three presentation lists, following a Latin-Square design, so that each participant saw 34 targets associated with each of the three prime types. Three additional lists were created by reversing the order of the items in each list, for a total of six presentation lists. The 102 experimental prime-target pairs were mixed with 438 prime-target filler pairs, for a total of 540 items in each list. In all prime-target filler pairs, the prime was an existing word. Of the filler targets, 270 were non-existing words, generated from existing English words using the software Wuggy (Keuleers & Brysbaert, 2010). A ‘no’ response was therefore required in 50% of the 540 trials. Furthermore, 102 of the filler pairs included an inflected *-ed* prime (51) or a derived *-er* prime (51) combined with a non-existing word target (e.g., *barked-LEAMS*). This way, the presentation of *-ed* and *-er* primes did not represent a cue for lexicality of the target. Finally, 68 of the filler targets were non-existing words that were orthographically embedded in their primes (e.g., *sincere-SINCH*), so that form overlap was also not a cue for a ‘yes’ response. Overall, 25.2% of the prime-target pairs in each presentation list were related, either morphologically or orthographically.

Procedure

Participants were tested in a quiet laboratory room. They were homogeneously assigned to one of the six lists. Participants’ accuracy and RTs in milliseconds were measured using the experimental software DMDX (Forster & Forster, 2003). We informed the participants that they would see a series of existing English words and invented words, and that they would have to indicate as quickly and as accurately as possible whether the word on

³ Lemma frequency here refers the dominant word class of the word. This means that, for words that can belong to different word classes (as is often the case in English; see e.g., ‘brew_{noun}’ and ‘brew_{verb}’), the word class that is considered for lemma frequency is the one most frequently associated to that word, in the Subtlex-UK corpus.

screen is an existing word ('lexical decision'). The 'yes' or 'no' responses were provided by pressing one of two buttons on a gamepad. Participants provided 'yes' responses with their dominant hand. Each trial started with a blank screen presented for 500 ms, followed by a mask consisting of ten hashes, also presented for 500 ms. Next, the prime appeared. The prime remained on screen for 50 ms and it was directly followed by the target. The target was displayed until a button press or until the timeout, which was set at 2000 ms. The next trial automatically began right after a button press or timeout. Primes were always presented in lowercase letters, and targets in uppercase letters.

Data analysis

Incorrect responses (7.47%) and timeouts (0.18%) were excluded from analysis. There were no other exclusions of participants, items, or data points. No short outliers were identified (the fastest response was 309 ms). RTs were analyzed with Bayesian mixed-effects distributional regression models, assuming a shifted-lognormal response distribution. Bayesian models combine prior information with the evidence from the data in order to obtain a probability distribution over a parameter's possible values—its *posterior distribution*. In this way, an experimental effect can be quantified in terms of the probability of its different magnitudes, which is more informative than a binary statement about whether an effect exists or not (McElreath, 2020; Vasishth et al., 2018). Prior distributions on all effects were centered around zero, with their width informed by domain knowledge of typical effect sizes in masked morphological priming (in L1 and L2). The chosen prior distributions ruled out effects that are extreme or unreasonable, but still allowed for a range of possible effects, both positive and negative (see Supplementary Materials S1, Table S1, for the full specification of prior distributions). Analyses were performed with the *brms* package in R (Bürkner, 2017; R Core Team, 2020). The procedures for fitting Bayesian models and assessing their convergence followed recent recommendations (Schad et al., 2020; Vasishth et al., 2018).

The statistical models included fixed effects for prime type (unrelated, inflected, derived) and trial position (centered). The prime type variable was coded with treatment contrasts, with the unrelated condition as the reference value. Thus, the models directly estimated the effects of inflectional and derivational priming, more specifically, the differences between responses in the inflected and derived conditions relative to the unrelated condition. The fixed effects of prime type and trial were estimated on both the *mu* and *sigma* parameters of the shifted-lognormal RT distribution (i.e., on the mean and standard deviation of RTs in the shifted log scale). Models included random intercepts for participant and target word (for both *mu* and *sigma*). Random slopes for prime type were not included, because model comparisons on the basis of ELPD (a Bayesian measure of predictive accuracy) showed that none of the random slopes (by participant and by item, for mean and for *sigma*) provided meaningful improvements in predictive accuracy, defined here as an ELPD difference larger than 2 standard errors (Bürkner, 2017; Vasishth et al., 2018). All models and code can be downloaded from <https://osf.io/4zwty>.

Results

Accuracy rates were very high in all conditions (unrelated: 90.15%; inflected: 93.73%; derived 93.69%). No further analyses of accuracy were conducted. Overall median RTs in

the three prime type conditions were 661 ms (unrelated), 625 ms (inflected), and 625 ms (derived).

A summary of the statistical model is shown in Table 2. For each estimate in Table 2, we report the mean of its posterior distribution together with its 95% credible interval (i.e., the range within which a parameter falls with 95% probability). With regards to the *mu* parameter (see top four rows in Table 2), both morphological primes were found to speed up responses (see negative coefficients). That is, mean RTs (in the shifted log scale) were shorter following inflected and derived primes than following unrelated primes (by 43 ms for inflected, and 39 ms for derived, in back-transformed estimates). These effects show relatively narrow 95% credible intervals, with upper bounds that are far away from zero, and thus indicate strong support for facilitation effects on the speed of lexical decision responses for both types of morphological primes. The effects of inflected and derived primes can also be directly compared; in a Bayesian framework, this is easily achieved by subtracting all pairs of posterior samples for one and the other effect. The resulting posterior distribution shows little evidence for a difference between the effects of the two prime types on *mu* (mean -0.011 95%CI [-0.031, 0.008]).

Table 2

Summary of a distributional model, with both mean and standard deviation (sigma) of (shifted) log-RTs predicted by prime type (unrelated, inflected, derived)

	Estimate	L-95% CI	U-95% CI
Intercept (unrelated)	5.937	5.878	5.998
Prime type (derived vs. unrelated)	-0.110	-0.128	-0.091
Prime type (inflected vs. unrelated)	-0.121	-0.140	-0.102
Trial (centered)	-0.020	-0.025	-0.015
Sigma ~ Intercept (unrelated)	0.314	0.297	0.332
Sigma ~ Prime type (derived vs. unrelated)	0.010	-0.004	0.024
Sigma ~ Prime type (inflected vs. unrelated)	0.025	0.010	0.040
Sigma ~ Trial (centered)	0.001	-0.002	0.005

Note. All estimates are in the shifted log-ms scale. The overall shift was estimated as 303 ms 95%CI: [298, 306].

The model also revealed effects of morphological primes on the *sigma* parameter, that is, on the standard deviation of RTs (again in the shifted log scale; see four bottom rows in Table 2). Specifically, the positive coefficients for both morphological priming conditions suggest an increase in variability relative to the unrelated condition. Figure 2, panels a and b, show the full posterior distributions for these effects. For derived primes (panel a), a large proportion of the posterior distribution is on the positive side (i.e., greater standard deviation in the derived than in the unrelated condition). However, there is still some mass on the negative side, and the 95% credible interval for this effect crosses zero, indicating that very small or even negative values are not completely implausible (Lindley, 1970; Rouder et al., 2018). There is much clearer support for an effect of inflected primes (again, relative to the unrelated condition), with almost all of its posterior distribution on the positive side (panel b). A visualization of the effects of prime type on *mu* and *sigma*, which shows the estimated

RT distributions in the different prime type conditions, is included in the Supplementary Materials S2 (Figure S1).

Critically, the effect of inflected and derived primes can also be directly compared; as above, we obtained a posterior distribution for the contrast between the two effects, which is shown in Figure 2, panel c. Although the difference between the two conditions is relatively small in magnitude, the exclusively positive 95% credible interval provides some evidence that inflected primes increased the standard deviation of (shifted log) RTs more than derived primes.⁴

⁴ We fitted additional models to address possible confounds of the priming effects on *sigma*. First, effects on *sigma* could arise from unmodelled between-participant variability (i.e., from the lack of random slopes) in priming effects on *mu*. Second, there are some differences between derived and inflected primes in their mean word form and lemma frequencies (see Table 1). However, models containing by-participant and by-item random slopes on *mu*, as well as models with frequency measures as covariates, revealed very similar effects to those reported in Figure 2. We conclude that the reported priming effects are not explainable by between-subject heterogeneity or by frequency differences. We thank Seamus Donnelly for these suggestions.

Figure 2

Posterior distributions for the effects of derived (panel a) and inflected primes (panel b) on the SDs of log-RTs, relative to unrelated primes, and for the comparison between inflected and derived primes (panel c). Each posterior distribution contains its mean and 95% credible interval displayed in numerical and graphical form, as well as the percentage of posterior samples that are on the positive side.

Discussion

Despite recommendations of looking beyond central tendencies of distributions in RT analyses (e.g., Heathcote et al., 1991), this rationale has hardly been applied to visual word recognition studies and it is virtually absent from morphological processing research (but see Balota et al., 2008; Hasenäcker et al., 2016; and Yap et al., 2006). In the present study, we tried to close this gap and investigated masked morphological priming effects both on mean RTs to target words and on RT variability.

Based on the previous masked priming literature involving L2 speakers (Jacob et al., 2018; Kirkici & Clahsen, 2013; Silva & Clahsen, 2008; Veríssimo et al., 2018), we predicted priming effects on *mean* lexical decision RTs from derived words, but reduced or absent priming from inflected forms. However, our results showed that both types of morphologically related primes sped up word recognition latencies (and to a similar extent). Although unexpected, the finding of a priming effect with inflected words in L2 speakers has been previously reported in at least two other masked priming studies (Feldman et al., 2010; Foote, 2017), though these did not include derived primes. Our key prediction concerned effects of morphologically related primes on the standard deviation (*sigma*) of log-RTs, which was our measure of variability in responses. Following the proposal that variability in RTs should reflect the ability to efficiently access word representations (e.g., N. S. Segalowitz, 2008), and based on the previous evidence suggesting persistent vulnerabilities in L2 speakers in reliably accessing inflected forms (Blom et al., 2006; Hopp, 2013; White, 2003), we predicted that the presentation of inflected primes would produce more inconsistent benefits as compared to derived primes, leading to more variable lexical decision times. We found that the presentation of both derived and inflected primes led to larger RT variability as compared to unrelated primes, but inflected primes indeed increased RT variability more so than derived primes.

Our results suggest that, at least in an L2, accessing ‘print’ given ‘printed’ works less automatically than accessing ‘print’ given ‘printer’. Note that, although reliance on orthographic cues possibly plays a prominent role in L2 morphological processing (e.g., Heyer & Clahsen, 2015; J. Li et al., 2017), this cannot explain our effect on RT variability, since the two morphological conditions were perfectly pairwise matched with regard to prime–target orthographic overlap. Therefore, the effect we report is more likely to result from specific differences between inflection and derivation. In spite of their superficial similarity, derived and inflected words modify their stems in different ways: while derivational operations create new lexical entries, inflectional operations are purely grammatical, in that they spell out morphosyntactic features of the stem, such as number, case, or tense (see Anderson, 1992). As mentioned in the Introduction, N. S. Segalowitz and Segalowitz (1993) and Solovyeva and DeKeyser (2018) have observed that variability increases in the process of establishing lexical representations, while it decreases when these are well established. We therefore suggest that the different lexical status of inflected and derived words is responsible for the differences in processing variability we observed: processing complex forms which have established lexical representations (in our case, derived words) and accessing the representations of their stems works more automatically, or efficiently, than processing complex forms that are bare spell-outs of grammatical properties of their stems (inflected forms). This should be particularly

true in L2 populations, given that morphosyntax remains a particularly vulnerable domain in L2 acquisition, and even proficient L2 speakers may not be able to consistently access the information contained in inflected forms (Blom et al., 2006; Hopp, 2013; White, 2003). In this way, the framework proposed by N. S. Segalowitz and Segalowitz (1993) and Solovyeva and DeKeyser (2018) can very well—and quite economically—account for our findings.

An open question is to what extent what we found also applies to native processing. Theoretically, the different nature of derived and inflected words may also affect the variability of responses in native speakers. Alternatively, the effects we have obtained may be specific to non-native or less proficient speakers, given their particularly vulnerability in the domains of inflection and morphosyntax. Future studies should examine morphological priming effects on RT variability across different populations, including native speakers and speakers at different proficiency levels. The detection of such effects in native speakers may nevertheless prove to be particularly challenging, considering that native speakers display higher language proficiency, less variable RTs, and show consistent and robust priming effects across all types of morphologically related primes, at least in the masked priming paradigm.

Accounts of morphological processing based on discriminative learning, which do not posit an independent level of morphological representation (e.g. Baayen et al., 2018; Baayen et al., 2011; Baayen & Smolka, 2020), are also compatible with our data. A key tenet of discriminative learning is that words with similar meanings are more difficult to discriminate from each other. Consequently, words of an inflectional paradigm should be more difficult to discriminate than derived words, as the former are characterized by high semantic similarity between each other. Accessing the meaning of inflected words should therefore lead to greater inconsistencies as compared to derived words, resulting in larger RT variability. This problem is likely to be particularly strong for L2 speakers, since they are faced with the near-synonymy of translation equivalents between their L1 and their L2 (see Chuang, Bell, et al., 2021). However, considering the small size of English inflectional paradigms, it would remain to be tested whether English-inflected words are indeed more difficult to discriminate than English-derived words, especially in the case of the very productive and transparent *-er* derivations that were tested in the present study.

A relevant aspect of the study concerns the type of RT distribution that we modelled. Similarly to the studies on RT variability reported above, our goal was to investigate variability in responses as measured separately from their central tendency. However, in RT distributions, the mean and standard deviation of the distribution have a linear relationship (Wagenmakers & Brown, 2007), such that manipulations that slow down responses (i.e., lead to larger mean RTs in milliseconds) also increase their variability (i.e., lead to larger SDs). By modeling a (shifted) lognormal distribution, we could estimate effects on the two parameters of interest, *mu* and *sigma*, independently of one another, thus being able to isolate effects on variability from bare speed-up effects. This is very well illustrated by the fact that derived and inflected primes had different effects on *sigma* (i.e., on the standard deviation of shifted log-RTs), although the two prime types were associated with very similar effects on *mu* (i.e., on the mean). Note, however, that the shifted-lognormal distribution is only one of the possible distributions that can be used to model RT data. Other theoretical distributions that can capture the properties of RT data are the gamma and the inverse-Gaussian distributions

(see Lo & Andrews, 2015). However, these are characterized by different parameters (e.g., the tail *tau*) and different relationships between mean and SD. Therefore, it is worth stressing that the effects we are reporting here pertain to the SD of a distribution of shifted log-RTs, and may not apply to other distributions.

To conclude, we have shown that morphological type (derivation vs. inflection) modulates the processing of complex words as tested with the masked priming paradigm, with inflected primes in particular producing larger increases in the variability of responses. By examining only the central tendency (i.e., the mean) of response times, we would have been unable to distinguish priming effects produced by derived and inflected words. However, the differences between the two prime types emerged when going beyond mean RTs by specifically investigating the variability of responses. Besides the relevance of the present results for models of morphological processing, the current study therefore emphasizes the importance of testing predictions on other parameters of the RT distribution rather than just its central tendency (Balota et al., 2008; Heathcote et al., 1991), and more generally, of treating variability in performance as a direct object of investigation in psycholinguistics.

Supplementary files

- S1. Prior distributions on each estimate of the Bayesian distributional model
- S2. Plot of estimated distributions in the different prime type conditions
- S3. Alternative analysis with generalized additive mixed models (GAMMs)

Open practices statement

The materials, the raw data, and the analysis scripts related to this manuscript are available at <https://osf.io/4zwty>.

References

- Allan, D. (2004). *Oxford Placement test 1*. Oxford University Press.
- Amenta, S., & Crepaldi, D. (2016). Editorial: The Variable Mind? How Apparently Inconsistent Effects Might Inform Model Building. *Frontiers in Psychology*, 7. <https://doi.org/10.3389/fpsyg.2016.00185>
- Amenta, S., & Crepaldi, D. (2012). Morphological processing as we know it: An analytical review of morphological effects in visual word identification. *Frontiers in Psychology*, 3. <https://doi.org/ggvfm5>
- Anderson, S. R. (1992). *A-Morphous morphology*. Cambridge University Press.
- Andrews, S. (2012). Individual differences in skilled visual word recognition and reading: The role of lexical quality. In *Visual word recognition: Meaning and context, individuals and development*, Vol. 2 (pp. 151–172). Psychology Press.
- Baayen, R. H. (2014). Experimental and Psycholinguistic Approaches. In R. Lieber & P. Štekauer (Eds.), *The Oxford Handbook of Derivational Morphology* (pp. 95–117).
- Baayen, R. H., Chuang, Y.-Y., & Blevins, J. P. (2018). Inflectional morphology with linear mappings. *The Mental Lexicon*, 13(2), 230–268. <https://doi.org/10.1075/ml.18010.baa>
- Baayen, R. H., Milin, P., Đurđević, D. F., Hendrix, P., & Marelli, M. (2011). An amorphous model for morphological processing in visual comprehension based on naive discriminative learning. *Psychological Review*, 118(3), 438. <https://doi.org/10.1037/a0023851>
- Baayen, R. H., & Smolka, E. (2020). Modeling Morphological Priming in German With Naive Discriminative Learning. *Frontiers in Communication*, 5. <https://doi.org/10.3389/fcomm.2020.00017>
- Baayen, R. H., Vasishth, S., Kliegl, R., & Bates, D. (2017). The cave of shadows: Addressing the human factor with generalized additive mixed models. *Journal of Memory and Language*, 94, 206–234. <https://doi.org/10.1016/j.jml.2016.11.006>
- Balota, D. A., & Yap, M. J. (2011). Moving beyond the mean in studies of mental chronometry: The power of response time distributional analyses. *Current Directions in Psychological Science*, 20(3), 160–166. <https://doi.org/bkcjq6>
- Balota, D. A., Yap, M. J., Cortese, M. J., & Watson, J. M. (2008). Beyond mean response latency: Response time distributional analyses of semantic priming. *Journal of Memory and Language*, 59(4), 495–523. <https://doi.org/fs3gpg>
- Bialystok, E., & Hakuta, K. (1999). Confounded age: Linguistic and cognitive factors in age differences for second language acquisition. In D. Birdsong (Ed.), *Second language acquisition and the critical period hypothesis*. Erlbaum.
- Blom, E., Polisšenská, D., & Weerman, F. (2006). Effects of age on the acquisition of agreement inflection. *Morphology*, 16(2), 313–336. <https://doi.org/c56gqf>
- Bürkner, P.-C. (2017). brms: An R package for Bayesian multilevel models using Stan. *Journal of Statistical Software*, 80(1), 1–28. <https://doi.org/gddxwp>
- Bürkner, P.-C. (2018). Advanced Bayesian multilevel modeling with the R package brms. *The R Journal*, 10(1), 395–411. <https://doi.org/gfxzpn>
- Chuang, Y.-Y., Bell, M. J., Banke, I., & Baayen, R. H. (2021). Bilingual and Multilingual Mental Lexicon: A Modeling Study With Linear Discriminative Learning. *Language Learning*, 71(S1), 219–292. <https://doi.org/10.1111/lang.12435>
- Chuang, Y.-Y., Fon, J., Papakyritsis, I., & Baayen, R. H. (2021). Analyzing Phonetic Data with Generalized Additive Mixed Models. In *Manual of Clinical Phonetics*. Routledge.

- Ciaccio, L. A., & Clahsen, H. (2020). Variability and Consistency in First and Second Language Processing: A Masked Morphological Priming Study on Prefixation and Suffixation. *Language Learning*, 70(1), 103–136. <https://doi.org/10.1111/lang.12370>
- Ciaccio, L. A., Kgolo, N., & Clahsen, H. (2020). Morphological decomposition in Bantu: A masked priming study on Setswana prefixation. *Language, Cognition and Neuroscience*, 1–15. <https://doi.org/10.1080/23273798.2020.1722847>
- Council of Europe. (2001). *Common European Framework of Reference for Languages: Learning, teaching, assessment*. Cambridge University Press.
- Diependaele, K., Duñabeitia, J. A., Morris, J., & Keuleers, E. (2011). Fast morphological effects in first and second language word recognition. *Journal of Memory and Language*, 64(4), 344–358. <https://doi.org/fjghsn>
- Feldman, L. B. (2000). Are morphological effects distinguishable from the effects of shared meaning and shared form? *Journal of Experimental Psychology: Learning, Memory, and Cognition*, 26(6), 1431–1444. <https://doi.org/http://dx.doi.org/10.1037/0278-7393.26.6.1431>
- Feldman, L. B., Kostić, A., Basnight-Brown, D. M., Đurđević, D. F., & Pastizzo, M. J. (2010). Morphological facilitation for regular and irregular verb formations in native and non-native speakers: Little evidence for two distinct mechanisms. *Bilingualism: Language and Cognition*, 13(02), 119. <https://doi.org/10.1017/S1366728909990459>
- Feldman, L. B., & Soltano, E. G. (1999). Morphological Priming: The Role of Prime Duration, Semantic Transparency, and Affix Position. *Brain and Language*, 68(1-2), 33–39. <https://doi.org/10.1006/brln.1999.2077>
- Foote, R. (2017). The storage and processing of morphologically complex words in L2 Spanish. *Studies in Second Language Acquisition*, 39(4), 735–767. <https://doi.org/10.1017/S0272263115000376>
- Forster, K. I., & Forster, J. C. (2003). DMDX: A Windows display program with millisecond accuracy. *Behavior Research Methods, Instruments, & Computers*, 35(1), 116–124. <https://doi.org/d5dfns>
- Grainger, J., & Beyersmann, E. (2017). Edge-Aligned Embedded Word Activation Initiates Morpho-Orthographic Segmentation. In B. H. Ross (Ed.), *Psychology of Learning and Motivation* (Vol. 67, pp. 285–317). Academic Press. <https://doi.org/10.1016/bs.plm.2017.03.009>
- Hasenäcker, J., Beyersmann, E., & Schroeder, S. (2016). Masked Morphological Priming in German-Speaking Adults and Children: Evidence from Response Time Distributions. *Frontiers in Psychology*, 7. <https://doi.org/10.3389/fpsyg.2016.00929>
- Heathcote, A., Popiel, S. J., & Mewhort, D. J. (1991). Analysis of response time distributions: An example using the Stroop task. *Psychological Bulletin*, 109(2), 340–347. <https://doi.org/10.1037/0033-2909.109.2.340>
- Heyer, V., & Clahsen, H. (2015). Late bilinguals see a *Scan* in *Scanner* AND in *Scandal* : Dissecting formal overlap from morphological priming in the processing of derived words. *Bilingualism: Language and Cognition*, 18(3), 543–550. <https://doi.org/10.1017/S1366728914000662>
- Hopp, H. (2013). Grammatical gender in adult L2 acquisition: Relations between lexical and syntactic variability. *Second Language Research*, 29(1), 33–56. <https://doi.org/f4rnbm>
- Jacob, G., Heyer, V., & Veríssimo, J. (2018). Aiming at the same target: A masked priming

- study directly comparing derivation and inflection in the second language. *International Journal of Bilingualism*, 22(6), 619–637. <https://doi.org/ggvfsr>
- Keuleers, E., & Brysbaert, M. (2010). Wuggy: A multilingual pseudoword generator. *Behavior Research Methods*, 42(3), 627–633. <https://doi.org/cfxkgw>
- Kidd, E., Donnelly, S., & Christiansen, M. H. (2018). Individual Differences in Language Acquisition and Processing. *Trends in Cognitive Sciences*, 22(2), 154–169. <https://doi.org/10.1016/j.tics.2017.11.006>
- Kinoshita, S., & Lupker, S. J. (Eds.). (2004). *Masked priming: The state of the art*. Psychology Press.
- Kirkici, B., & Clahsen, H. (2013). Inflection and derivation in native and non-native language processing: Masked priming experiments on Turkish. *Bilingualism: Language and Cognition*, 16(4), 776–791. <https://doi.org/ggvjsg>
- Kneib, T., & Umlauf, N. (2017). *A primer on Bayesian distributional regression* (Working Papers in Economics and Statistics No. 2017-13). University of Innsbruck, Research Platform Empirical and Experimental Economics (eeecon).
- Lemhöfer, K., & Broersma, M. (2012). Introducing LexTALE: A quick and valid Lexical Test for Advanced Learners of English. *Behavior Research Methods*, 44(2), 325–343. <https://doi.org/c9f897>
- Li, J., Taft, M., & Xu, J. (2017). The Processing of English Derived Words by Chinese-English Bilinguals. *Language Learning*, 67(4), 858–884. <https://doi.org/10.1111/lang.12247>
- Li, M., Jiang, N., & Gor, K. (2017). L1 and L2 processing of compound words: Evidence from masked priming experiments in English. *Bilingualism: Language and Cognition*, 20(02), 384–402. <https://doi.org/10.1017/S1366728915000681>
- Lindley, D. V. (1970). *Introduction to probability and statistics from a Bayesian viewpoint. Part 2: inference* (2nd ed.). Cambridge University Press.
- Lo, S., & Andrews, S. (2015). To transform or not to transform: Using generalized linear mixed models to analyse reaction time data. *Frontiers in Psychology*, 6. <https://doi.org/10.3389/fpsyg.2015.01171>
- Logan, G. D. (1992). Shapes of reaction-time distributions and shapes of learning curves: A test of the instance theory of automaticity. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, 18(5), 883–914. <https://doi.org/bwv74n>
- Matzke, D., & Wagenmakers, E.-J. (2009). Psychological interpretation of the ex-Gaussian and shifted Wald parameters: A diffusion model analysis. *Psychonomic Bulletin & Review*, 16(5), 798–817. <https://doi.org/10.3758/PBR.16.5.798>
- McElreath, R. (2020). *Statistical rethinking: A Bayesian course with examples in R and Stan* (Second). CRC Press.
- Nicenboim, B., Logachev, P., Gattei, C., & Vasishth, S. (2016). When high-capacity readers slow down and low-capacity readers speed up: Working memory and locality effects. *Frontiers in Psychology*, 7. <https://doi.org/ggks56>
- R Core Team. (2020). *R: A language and environment for statistical computing*. R Foundation for Statistical Computing.
- Rastle, K., Davis, M. H., Marslen-Wilson, W. D., & Tyler, L. K. (2000). Morphological and semantic effects in visual word recognition: A time-course study. *Language and Cognitive Processes*, 15(4-5), 507–537. <https://doi.org/10.1080/01690960050119689>
- Rastle, K., Davis, M. H., & New, B. (2004). The broth in my brother's brothel: Morpho-

- orthographic segmentation in visual word recognition. *Psychonomic Bulletin & Review*, 11(6), 1090–1098. <https://doi.org/ddq274>
- Rodgers, D. M. (2011). The automatization of verbal morphology in instructed second language acquisition. *IRAL - International Review of Applied Linguistics in Language Teaching*, 49(4). <https://doi.org/10.1515/iral.2011.016>
- Rouder, J. N. (2005). Are unshifted distributional models appropriate for response time? *Psychometrika*, 70(2), 377–381. <https://doi.org/c9cx9s>
- Rouder, J. N., Haaf, J. M., & Vandekerckhove, J. (2018). Bayesian inference for psychology, part IV: Parameter estimation and Bayes factors. *Psychonomic Bulletin & Review*, 25(1), 102–113. <https://doi.org/gc9qfx>
- Schad, D. J., Betancourt, M., & Vasishth, S. (2020). Toward a principled Bayesian workflow in cognitive science. *Psychological Methods*. <https://doi.org/10.1037/met0000275>
- Segalowitz, N. S. (2008). Automaticity and Second Languages. In C. J. Doughty & M. H. Long (Eds.), *The handbook of second language acquisition* (Vol. 27, pp. 382–408). Blackwell.
- Segalowitz, N. S., & Segalowitz, S. J. (1993). Skilled performance, practice, and the differentiation of speed-up from automatization effects: Evidence from second language word recognition. *Applied Psycholinguistics*, 14(3), 369–385. <https://doi.org/10.1017/S0142716400010845>
- Segalowitz, S. J., Segalowitz, N. S., & Wood, A. G. (1998). Assessing the development of automaticity in second language word recognition. *Applied Psycholinguistics*, 19(1), 53–67. <https://doi.org/10.1017/S0142716400010572>
- Silva, R., & Clahsen, H. (2008). Morphologically complex words in L1 and L2 processing: Evidence from masked priming experiments in English. *Bilingualism: Language & Cognition*, 11, 245–260. <https://doi.org/10.1017/S1366728908003404>
- Solovyeva, K., & DeKeyser, R. (2018). Response time variability signatures of novel word learning. *Studies in Second Language Acquisition*, 40(1), 225–239. <https://doi.org/10.1017/S0272263117000043>
- Taft, M., & Forster, K. I. (1975). Lexical storage and retrieval of prefixed words. *Journal of Verbal Learning and Verbal Behavior*, 14(6), 638–647. [https://doi.org/10.1016/S0022-5371\(75\)80051-X](https://doi.org/10.1016/S0022-5371(75)80051-X)
- Tomaschek, F., Arnold, D., Sering, K., Tucker, B. V., van Rij, J., & Ramscar, M. (2020). Articulatory Variability is Reduced by Repetition and Predictability. *Language and Speech*, 0023830920948552. <https://doi.org/10.1177/0023830920948552>
- van Heuven, W. J. B., Mandera, P., Keuleers, E., & Brysbaert, M. (2014). Subtlex-UK: A New and Improved Word Frequency Database for British English. *Quarterly Journal of Experimental Psychology*, 67(6), 1176–1190. <https://doi.org/ggrw22>
- Vasishth, S., Nicenboim, B., Beckman, M. E., Li, F., & Kong, E. J. (2018). Bayesian data analysis in the phonetic sciences: A tutorial introduction. *Journal of Phonetics*, 71, 147–161. <https://doi.org/gfzq3c>
- Veríssimo, J., Heyer, V., Jacob, G., & Clahsen, H. (2018). Selective effects of age of acquisition on morphological priming: Evidence for a sensitive period. *Language Acquisition*, 25(3), 315–326. <https://doi.org/ggffzk>
- Wagenmakers, E.-J., & Brown, S. (2007). On the linear relation between the mean and the standard deviation of a response time distribution. *Psychological Review*, 114(3), 830–841. <https://doi.org/btnjz6>

- White, L. (2003). *Second language acquisition and Universal Grammar* (First). Cambridge University Press. <https://doi.org/10.1017/CBO9780511815065>
- Wood, S. N. (2017). GAMs in Practice: mgcv. In *Generalized Additive Models* (Second). Chapman and Hall/CRC Press.
- Yap, M. J., Balota, D. A., Cortese, M. J., & Watson, J. M. (2006). Single- versus dual-process models of lexical decision performance: Insights from response time distributional analysis. *Journal of Experimental Psychology: Human Perception and Performance*, 32(6), 1324–1344. <https://doi.org/10.1037/0096-1523.32.6.1324>

Supplementary Materials to:
Investigating variability in morphological processing with Bayesian distributional models

S1. Prior distributions on each estimate of the Bayesian distributional model

Table S1: Prior distributions on each estimate of the Bayesian distributional model

Parameter	Coefficient	Type	Prior	Units
Mu (mean)	Intercept	Fixed effect	Normal(6.0, 0.5)	log-ms
Mu (mean)	Intercept	Random effect	Normal(0, 0.5)	log-ms
Mu (mean)	Slopes	Fixed effect	Normal(0, 0.25)	log-ms
Sigma (SD)	Intercept	Fixed effect	Normal(-2, 1)	log of log-ms
Sigma (SD)	Intercept	Random effect	Normal(0, 1)	log of log-ms
Sigma (SD)	Slopes	Fixed effect	Normal(0, 0.5)	log of log-ms
Shift			Normal(250, 100)	ms

Note. Prior distributions are in the notation Normal(mean, SD). These are expressed in different units due to the parametrizations used by the brms R-package: estimates on mu (mean) are in log-ms because a (shifted-)lognormal response distribution was assumed; estimates on sigma (SD) are modelled in the log-scale by default because they must be strictly positive (hence, in log of log-ms). Note that all sigma effects reported in the paper were back-transformed to log-ms, for easier interpretation and for consistency with the other estimates.

S2. Plot of estimated distributions in the different prime type conditions

Figure S1 below shows the estimated RT distributions in the three prime type conditions (unrelated, inflected, derived), with estimates obtained from the Bayesian distributional model reported in the main paper. RTs are shown in the modelled shifted log-millisecond scale; thus, they are normally distributed with estimated μ and σ parameters. Plotting in a log scale allows bypassing the dependency between mean and standard deviation that exists for millisecond RTs, so that the effects of prime type can be visualised independently.

Note that, in a Bayesian framework, every quantity (i.e., every estimated parameter and every prediction) is associated with a full posterior distribution. Visualisation of distributions can thus become quite complex, as well as computationally intensive. The visualisation below is a simplification: it does not show the uncertainty associated with the predictions, but depicts only the idealised normal distributions that were estimated by the model, on the basis of point estimates (i.e., the mean of the posteriors) of both μ and σ , for each of the three conditions.

As can be seen in Fig. S1, prime type had effects on both μ and σ . The model estimated shorter mean RTs in the inflected and derived conditions than in the unrelated condition. With regards to σ , the effects were small, but can be seen by inspecting the heights of the three distributions: responses in the derived condition were estimated as slightly more variable than in the unrelated condition ('shorter' distribution), and those in the inflected condition were more variable than both.

Figure S1: Estimated distributions of RTs in the three prime type conditions (unrelated, inflected, derived). RTs are shown in the shifted log millisecond scale.

S3. Alternative analyses with Generalized Additive Mixed Models (GAMMs)

The present supplementary materials illustrate an additional way of analyzing reaction time (RT) distributions by making use of Generalized Additive Mixed Models (GAMMs); see Baayen et al. (2017)¹. Similarly to the Bayesian distributional model reported in the main paper (Ciaccio & Veríssimo, 2022), GAMMs allow to test for experimental effects on both the mean of (transformed) RTs, i.e. the *mu* parameter of the distribution, and their standard deviation (SD), i.e. the *sigma* parameter. This can be achieved by specifying `family = gaulss`. This way, we fit a Gaussian model containing two formulae, the first specifying the predictors for estimating effects on mean RTs, and the second the predictors for effects on their SD.

All analyses were performed using the `mgcv` package (Wood, 2017) version 1.8.40 in R version 4.2.0 (2022-04-22) (R Core Team, 2020). All models and code can be downloaded from <https://osf.io/4zwty>.

We first load the data and prepare the dataset for the analyses.

```
c. <- function(x) scale(as.numeric(x, scale=F)) # Function used to center variables

# Load data
varmorph <- read.csv("data/varmorph.csv", sep=";", header=T) # Read data

varmorph$set <- as.factor(varmorph$set)
varmorph$subj_id <- as.factor(varmorph$subj_id)
varmorph$target <- as.factor(varmorph$target)
varmorph$prime_type <- as.factor(varmorph$prime_type)

c_trialN <- c.(I(varmorph$trial_n/100)) # Center trial number

# - Keep only experimental items (exclude fillers)
# - Exclude incorrect responses and timeouts
# - Set unrelated prime as baseline

varmorph <- droplevels(varmorph[varmorph$set=="experimental",])
varmorph <- droplevels(varmorph[varmorph$accuracy==1,])
varmorph$prime_type <- relevel(varmorph$prime_type, "unrelated")
```

¹We thank Harald Baayen for suggesting this to us and providing useful instructions.

Before fitting our Gaussian model, we need to transform our RT data with an appropriate transformation. The transformation that best corresponds to the shifted log-normal distribution that we fitted in the main paper consists in log-transforming RTs after subtracting the ‘shift’, i.e. after shifting the entire distribution by an amount of milliseconds under which RTs are considered implausible. In our case, the ‘amount of ms’ we take is the shift estimated by the Bayesian distributional model reported in the main paper, i.e. 302.58 ms. As we can see from the plots below, shifted log-transforming our data (left panel) makes them closer to a Gaussian distribution than a simple log-transformation (middle panel).

Admittedly, taking a shift estimated by another Bayesian distributional model may make the analysis unnecessarily complex. At least for this dataset, the inverse transformation seems to also work quite well (see right panel below). For those analyzing data directly with GAMMs, this may be a more practical option, though probably at the expense of interpretability of the estimates. In this document, we will stick to the shifted log transformation for consistency with the main paper.

```
# log transformation
varmorph$RTlog <- log(varmorph$rt)

# shifted log-transformation, with estimated shift from Bayesian model (see main paper)
varmorph$RTslog <- log(varmorph$rt-302.58)

# inverse transformation
varmorph$RTinv <- -1000/varmorph$rt

par(mfrow=c(1,3))

qqnorm(varmorph$RTlog) # log transformation does not work so well (left panel)
qqline(varmorph$RTlog)

qqnorm(varmorph$RTslog) # shifted log transformation is better (middle panel)
qqline(varmorph$RTslog)

qqnorm(varmorph$RTinv) # inverse reaction times are also fine (right panel)
qqline(varmorph$RTinv)
```

In the following, we will start from a simple GAMM and make it successively more complex, leading to a final model with a very similar structure to the one reported in the main manuscript. We will additionally show that: (a) the final model testing effects on both the μ and the σ parameters has a better fit than one only testing effects on μ ; (b) the final model shows parallel results to the one reported in the manuscript regarding the main contrast we focused on, namely the difference between inflection and derivation with regards to the effect on σ .

We first fit a baseline mixed-effects model (m_0) testing only effects of Prime Type (inflected, derived, unrelated) on μ , i.e. on mean (shifted log) reaction times. The model contains by-subject and by-item random effects. It additionally includes the predictor (centered) Trial Number. Unlike in the main paper, because GAMMs allow for smooth predictors and the effect of trial number over the experiment is likely to be non-linear, we specify Trial Number as a smooth term by using `s()`. Starting from this model, we will now fit more complex models and evaluate whether they provide a better fit than the simpler model based on their REML score (lower REML = better fit).

```
m0 <- gam(RTslog ~ prime_type + s(c_trialN) +
          s(subj_id, bs="re") +
          s(target, bs="re"),
          data=varmorph, method="REML")
```

```
m0$gcv.ubre[1] # Get the REML
```

```
##      REML
## 2432.698
```

We next add the σ parameter to our model (m_1), and additionally add (centered) Trial Number as smooth predictor to the σ part of the formula (m_2). Note that, for consistency with the model reported in the main paper, we will keep m_2 although Trial Number does not improve the fit of m_1 . Therefore, at this step, the model fit after adding σ is slightly worse than the fit of m_0 . However, we haven't included random effects for the σ parameter yet.

```
m1 <- gam(list(RTslog ~ prime_type + s(c_trialN) +
              s(subj_id, bs="re") +
              s(target, bs="re"),
              ~ prime_type),
          data=varmorph, family="gauss", method="REML")
```

```
m2 <- gam(list(RTslog ~ prime_type + s(c_trialN) +
              s(subj_id, bs="re") +
              s(target, bs="re"),
              ~ prime_type + s(c_trialN)),
          data=varmorph, family="gauss", method="REML")
```

```
# Get the REML of m1 (left) and m2 (right)
c(m1$gcv.ubre[1], m2$gcv.ubre[1])
```

```
##      REML      REML
## 2431.954 2433.578
```


Adding by-item random effects (m3) improves model fit compared to all previous models, including those not containing the `sigma` parameter. Therefore, the presence of the `sigma` parameter seems to be justified.

```
m3 <- gam(list(RTslog ~ prime_type + s(c_trialN) +  
              s(subj_id, bs="re") +  
              s(target, bs="re"),  
              ~ prime_type + s(c_trialN) +  
              s(target, bs="re")),  
          data=varmorph, family="gauss", method="REML")
```

```
m3$gcv.ubre[1] # lower REML than m2 and m0
```

```
##      REML  
## 2391.449
```

Adding by-subject random effects (m4) further improves the model fit. Adding by-subject or by-item random slopes did not further improve the model fit (models including random slopes can be downloaded at <https://osf.io/4zwty>). Therefore, m4 is our final model. This has a similar structure to the one reported in the main manuscript. It provides a better fit than a model which only tests effects on μ (cfr. REML m4 = 2294.1, see summary below; vs. REML m0 = 2432.698). The residuals also seem fairly normal.

```
m4 <- gam(list(RTslog ~ prime_type + s(c_trialN) +  
              s(subj_id, bs="re") +  
              s(target, bs="re"),  
              ~ prime_type + s(c_trialN) +  
              s(subj_id, bs="re")+  
              s(target, bs="re")),  
          data=varmorph, family="gauss", method="REML")
```

```
qqnorm(resid(m4))  
qqline(resid(m4))
```

Printing the summary of the model, we can now evaluate its output. Note that, for the *mu* parameter, model estimates are provided in the same scale as the data to which we fitted the model (i.e., in the log scale). Instead, for the *sigma* parameter, they are in the log of log-ms. Concerning the effects of 'Prime Type' on *mu*, the model output shows similarly large priming effects (i.e., speed-ups in RTs) on mean RTs for inflected and derived primes (respectively, $\beta = -0.1212462$ and $\beta = -0.1098143$). Regarding effects on *sigma*, while both inflected and derived primes *increase* RT variability, this effect seems to be robust only for inflected primes, and it is larger for inflected primes than for derived primes (respectively, $\beta = 0.0810701$ and $\beta = 0.0326677$).

```
summary(m4)
```

```
##
## Family: gaussian
## Link function: identity logb
##
## Formula:
## RTslog ~ prime_type + s(c_trialN) + s(subj_id, bs = "re") + s(target,
##      bs = "re")
## ~prime_type + s(c_trialN) + s(subj_id, bs = "re") + s(target,
##      bs = "re")
##
## Parametric coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)      5.935188   0.028898 205.385 < 2e-16 ***
## prime_typederived -0.109814   0.009369 -11.720 < 2e-16 ***
## prime_typeinflected -0.121246   0.009592 -12.640 < 2e-16 ***
## (Intercept).1     -1.215314   0.028443 -42.728 < 2e-16 ***
## prime_typederived.1  0.032668   0.023323  1.401 0.161313
## prime_typeinflected.1 0.081070   0.023347  3.472 0.000516 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##              edf   Ref.df   Chi.sq p-value
## s(c_trialN)    1.001    1.001   61.609 <2e-16 ***
## s(subj_id)     66.228   68.000 2997.768 <2e-16 ***
## s(target)      92.963  101.000 1495.619 <2e-16 ***
## s.1(c_trialN)  1.002    1.002    0.546  0.461
## s.1(subj_id)   55.626   68.000  468.930 <2e-16 ***
## s.1(target)    62.877  101.000  296.472 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Deviance explained = 40.3%
## -REML = 2294.1 Scale est. = 1          n = 6512
```

Finally, by releveling the baseline to derived, we can extract the model estimates for the direct comparison between derived and inflected Prime Type. Inflected and derived primes reliably differ for their effect on σ , with inflected primes increasing RT variability more than derived primes, as reported in the main paper.

```
varmorph$prime_type <- relevel(varmorph$prime_type, "derived")
m4der <- update(m4)
```

```
summary(m4der)$p.table
```

##	Estimate	Std. Error	z value	Pr(> z)
## (Intercept)	5.82537407	0.028908640	201.509796	0.000000e+00
## prime_typeunrelated	0.10981431	0.009369494	11.720410	1.001859e-31
## prime_typeinflected	-0.01143185	0.009621812	-1.188118	2.347870e-01
## (Intercept).1	-1.18264655	0.028222184	-41.904856	0.000000e+00
## prime_typeunrelated.1	-0.03266774	0.023322960	-1.400669	1.613131e-01
## prime_typeinflected.1	0.04840240	0.023063800	2.098631	3.584944e-02

References

- Baayen, R. H., Vasishth, S., Kliegl, R., & Bates, D. (2017). The cave of shadows: Addressing the human factor with generalized additive mixed models. *Journal of Memory and Language*, 94, 206–234. <https://doi.org/10.1016/j.jml.2016.11.006>
- Ciaccio, L. A., & Veríssimo, J. (2022). Investigating variability in morphological processing with Bayesian distributional models. *Psychonomic Bulletin & Review*, Advance online publication. <https://doi.org/10.3758/s13423-022-02109-w>
- R Core Team. (2020). *R: A language and environment for statistical computing*. R Foundation for Statistical Computing.
- Wood, S. N. (2017). GAMs in Practice: mgcv. In *Generalized Additive Models* (2nd ed.). Chapman and Hall/CRC Press.